Appendix: Confidence Limits and Point Estimation of Several Regression Lines with a Common $X$-intercept.
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Conventional methods of regression analysis provide formulas for estiates and confidence limits for the Y-intercept $\alpha$ and the slope $\beta$ of the linear regression of $Y$ on $X$. In the present circumstance, where $Y$ is water inflow and $X$ is osmotic pressure difference, a quantity of intrinsic interest is the $X$ intercept $M=-\alpha / \beta$ representing the osmotic pressure difference which results in zero average inflow of water. The X-intercept is likewise a basic paraneter in bioassay analysis and in that context the following formula has been developed (Finney - ) for 95 percent confidence limits on $M$ :

$$
\left[m-k^{2} \bar{X} \pm k \sqrt{(\bar{X}-m)^{2}+\left(1-k^{2}\right) \sum(X-\bar{X})^{2} n^{-1}}\right] /\left(1-k^{2}\right)
$$

The variables appearing in this formula are functions of the estimated slope and Y-intercept,

$$
\mathrm{b}=\frac{\left.\sum_{\sum_{i}\left(X_{i}-\bar{X}\right)\left(Y_{i}-\bar{Y}\right)}^{n} \quad a=\bar{Y}-b \bar{X}, \text {, } X_{i}-\bar{X}\right)^{2}}{l} \quad a=
$$

the standard error of $b$,

$$
s_{b}=\sqrt{\frac{\sum\left(Y_{i}-a-b X_{i}\right)^{2}}{\frac{n}{(n-2) \Sigma\left(X_{i}-\bar{X}\right)^{2}}}}
$$

and the value of Students $t$ at the 5 percent level on $n-2$ degrees of freedom; thus,
ts

$$
m=-\frac{a}{b} \quad k=\frac{b}{b}
$$

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In the present case where several different experimental conditions might be expectea to result in the same intercept for zero water uptake, a method is required for combining data from several experiments with different slopes and Y-intercepts in order to estimate and set confidence limits on their common X-intercept. If the residual variances

$$
s_{i}^{2}=\sum_{j=1}^{n_{i}}\left(Y_{i j}-a_{i}-b_{i} X_{i j}\right)^{2} /\left(n_{i}-2\right)
$$

are homogeneous for the several experiments then the following formula gives the Jesirea confỉence limits:

$$
\left[m-K^{2} \tilde{X}+K \sqrt{(\tilde{X}-m)^{2}+\left(1-K^{2}\right)\left(C_{X}-\tilde{X}^{2}\right)}\right] /\left(1-K^{2}\right) .
$$

In this case $m$ is the negative ratio of the sum of $Y$-intercepts and the sum of slopes,

$$
m=\stackrel{r}{-\sum a_{i}} \underset{l}{ } \underset{l}{r} \mathrm{~b}_{\mathrm{i}} .
$$

The quantity $\tilde{X}$ is a weighted average or the $r$ experimental means $\bar{X}_{i}$, with weights given by

$$
\frac{1}{w_{i}}=\sum_{j=1}^{n}\left(x_{i j}-\bar{x}_{i}\right)^{2}
$$

so that
and

$$
C_{X}=\stackrel{r}{\sum_{I}\left(\frac{1}{n_{i}}+w_{i} \bar{X}_{i}^{2}\right) / \stackrel{r}{\sum_{l} w_{i}} . . . . ~}
$$

The pooled residual variance is

$$
s^{2}=\sum_{1}^{r}\left(n_{i}-2\right) s_{i}^{2} / \Sigma\left(n_{i}-2\right)
$$

and $K$ is then defined by

$$
K=\frac{t s \sqrt{\Sigma \mathrm{w}}}{\Sigma \mathrm{~b}}
$$

where $t$ is now based on $\Sigma(n-2)$ degrees of freed cm.
This analysis is illustrated in detail below, utilizing the data from Experiments 1 and 2 as displayed in Figure 4.

Experiment I

| Observation | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| X | -75 | -67 | -36 | -32 | 4 | 6 | 32 | 41 | 58 | 77 |
| Y | -27 | -20 | -14 | -17 | -3 | 0 | 6 | 10 | 22 | 22 |

$$
\begin{aligned}
& n_{1}=10 \\
& \Sigma X_{1}=8 \\
& \Sigma X_{1}^{2}=24484 \\
& \Sigma X_{1} Y_{1}=7973 \\
& \frac{1}{10}\left(\Sigma X_{1}\right)^{2}=6.4 \\
& \frac{1}{10}\left(\Sigma X_{1}\right)\left(\Sigma Y_{1}\right)=16.8 \\
& \Sigma \overline{\left(X_{1}-\bar{X}_{1}\right)^{2}=24477.6} \\
& \Sigma\left(X_{1}-\bar{X}_{1}\right)\left(Y_{1}-\bar{Y}_{1}\right)=7989.8 \\
& \Sigma Y_{I}=-21 \\
& \Sigma \mathrm{Y}_{1}^{2}=2727 \\
& \frac{1}{10}\left(\Sigma Y_{1}\right)^{2}=44.1 \\
& \Sigma\left(Y_{1}-\bar{Y}_{1}\right)^{2}=2682.9 \\
& w_{1}=1 / 24477.6=.00004085 \quad b_{1}=7989.8 w_{1}=.3264 \quad \alpha_{1}=\frac{-21-8 b_{1}}{10}=-2.36 \\
& s_{1}^{2}=\frac{1}{8}\left(2682.9-(7989.8)^{2} w_{1}\right)=9.36 \quad s_{b_{1}}=\sqrt{w_{1} s_{1}^{2}}=.0196 \\
& m_{1}=\frac{2.36}{.3264}=7.2, \quad k_{1}=\frac{2.306(.0196)}{.3264}=.1385 \quad k_{1}^{2}=.0192
\end{aligned}
$$

$$
\frac{7.23-.0192(.8)+.1385 \sqrt{(.8-7.23)^{2}+(1-.0192)(2447.76)}}{1-.0192}=7.36 \pm 6.98
$$

## Experiment 2



$$
\begin{aligned}
& \Sigma Y_{2}=-4 \\
& \Sigma Y_{2}^{2}=898 \\
& \frac{1}{10}\left(\Sigma Y_{2}\right)^{2}=1.6 \\
& \overline{\Sigma\left(\mathrm{Y}_{2}-\overline{\mathrm{Y}}_{2}\right)^{2}=896.4} \\
& w_{2}=1 / 28030.9=.00003567 \quad b_{2}=4851.4 w_{2}=.1731 \quad a_{2}=\frac{1}{10}\left(-4-21 b_{2}\right)=-.7 t \\
& s_{2}^{2}=\frac{1}{8}\left(896.4-(4851.4)^{2}{w_{1}}_{1}\right)=7.10 \quad s_{b_{2}}=\sqrt{w_{2} s_{2}^{2}}=.0159 \\
& m_{2}=\frac{.76}{.1731}=4.39 \quad k_{1}=\frac{2.306(.0159)}{.1731}=.2118 \quad k_{2}^{2}=.0449 \\
& \frac{4.39-.0449(2.1) \pm .2118 \sqrt{(2.1-4.39)^{2}+(1-.0449)(2803.09)}}{1-.0449}=4.50 \pm 11 .+6
\end{aligned}
$$

$$
\begin{aligned}
& \mathrm{m}=\frac{2.36+.76}{.3264+.1731}=6.24 \quad X=\frac{4085(.8)+3567(2.1)}{4085+3567}=1.41 \\
& C_{X}=\frac{.1+.1+.00004085(.8)^{2}+.00003567(2.1)^{2}}{.00004085+.00003567}=2616.09 \\
& C_{X}-\tilde{X}^{2}=2614.10 \quad s^{2}=\frac{8(9.36)+8(7.10)}{8+8}=8.23 \quad \sqrt{s^{2} \Sigma \mathrm{w}}=.0251 \\
& K=\frac{2.120(.0251)}{.3264+.1731}=.1065 \quad K^{2}=.0113 \\
& \frac{\left.6.24-.0113(1.41) \pm .1065 \sqrt{(1.41-6.24)^{2}+(1-.0113)(2614.10}\right)}{1-.0113}
\end{aligned}
$$

While the preceding calculations provide valid confidence limits for the common $X$-intercept $M$, neither the estimate $m=6.24$ nor the confidence interval midpoint 6.30 represents the most efficient point estimate of $M$. The preferred estimates both of $M$ and of the slopes $\beta_{2}$ and $\beta_{2}$ of the two regression lines intersecting at $M$ are the maximum likelihood estimates obtained as the solution to the equations:
$\hat{M} \frac{\sum_{i=1}^{r} n_{i} \hat{\beta}_{i}{ }^{2} \bar{X}_{i}-\sum_{i}^{r} n_{i} \hat{\beta}_{i} \bar{Y}_{i}}{\sum_{i=1}^{r} n_{i} \hat{\beta}_{i}^{2}}$

$$
\hat{\beta}_{i}=\frac{\sum_{j=1}^{n_{i}} X_{i j} Y_{i j}-\hat{M}{ }_{1}^{\sum_{1}^{i}} Y_{i j}}{\sum_{j=1}^{n_{i}}\left(X_{i j}-\bar{X}_{i}\right)^{2}+n_{i}\left(\bar{X}_{i}-\hat{M}\right)^{2}}
$$

The iterative solution to these equations is illustrated below with the data from Experiments 1 and 2. Initial values to start the iteration are taken as $\hat{\beta}_{1}=b_{1}$ and $\hat{\beta}_{2}=b_{2}$, giving the initial value for $\hat{M}$ as:
$\hat{M}_{(0)}=\frac{(.3264)^{2}(8)+(.1731)^{2}(21)-(.3264)(-21)-(.1731)(-4)}{10(.3264)^{2}+10(.1731)^{2}}=6.61$

New values for $\hat{S}_{1}$ and $\hat{B}_{2}$ are then obtained from $\hat{\mathrm{M}}_{(0)}$ as:

$$
\hat{\beta}_{1}=\frac{7973-6.61(-21)}{24477.6+10(.8-6.61)^{2}}=.32689 \quad \hat{\beta}_{z}=\frac{-6+j-6.61(-4)}{28030.9+10(2.1-6.61)^{2}}=.17247
$$

The next trial value for $\hat{k}$ is then

$$
\hat{M}_{(1)}=\frac{(.32689)^{2}+(.17247)^{2}(21)-(.32609)(-21)-(.17247)(-4)}{10(.32689)^{2}+10(.17247)^{2}}=6.01 \overline{ }
$$

Evidently, further iterations would require the use of more decimal places than are needed, so the process stops here.

