

Time to Extinction of a Pure Death Process

by

D. L. Solomon

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Abstract

The expected time to extinction of a homogeneous pure death process is calculated in two ways. The fact that the two results must be equal establishes a proof of the

identity 
$$\sum_{i=1}^n \frac{1}{i} = \sum_{i=1}^n (-1)^{i-1} \binom{n}{i} \frac{1}{i} .$$

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Suppose that  $\{N(t); t \geq 0\}$  is a homogeneous pure death process with constant force of mortality  $\mu$  and initial size  $N(0) = N_0$ . Then for  $n = 0, 1, 2, \dots, N_0$  and  $t \geq 0$ ,

$$P[N(t) = n] = \binom{N_0}{n} e^{-\mu n t} (1 - e^{-\mu t})^{N_0 - n}.$$

Let  $T$  be the time of extinction of this process. Then the c.d.f. of  $T$  is

$$G_T(t) = P[T \leq t] = P[N(t) = 0] = (1 - e^{-\mu t})^{N_0}, \quad t \geq 0$$

with corresponding p.d.f.

$$g_T(t) = \frac{d}{dt} G_T(t) = N_0 \mu e^{-\mu t} (1 - e^{-\mu t})^{N_0 - 1}, \quad t \geq 0.$$

$$\text{Now } ET = N_0 \mu \int_0^{\infty} t e^{-\mu t} (1 - e^{-\mu t})^{N_0 - 1} dt$$

$$= N_0 \mu \sum_{i=0}^{N_0 - 1} \binom{N_0 - 1}{i} (-1)^i \int_0^{\infty} t e^{-\mu t (i+1)} dt$$

$$= N_0 \mu \sum_{i=0}^{N_0 - 1} \binom{N_0 - 1}{i} (-1)^i \frac{1}{\mu^2 (i+1)^2} = \frac{N_0}{\mu} \sum_{i=0}^{N_0 - 1} \frac{\binom{N_0 - 1}{i}}{(i+1)^2} (-1)^i$$

$$= \frac{1}{\mu} \sum_{i=1}^{N_0} (-1)^i \binom{N_0}{i} \frac{1}{i}.$$

Alternatively, let  $T_i$  be the time to first death in a population of initial size  $i$  undergoing a pure death process with parameter  $\mu$ . Then  $T_i$  has c.d.f.

$$\begin{aligned} F_i(t) &\equiv P[T_i \leq t] = 1 - P[T_i > t] = 1 - P[N(t) = i | N(0) = i] \\ &= 1 - e^{-i\mu t}, \quad t \geq 0, \end{aligned}$$

with corresponding density

$$f_i(t) = i\mu e^{-i\mu t}, \quad t \geq 0 \text{ and}$$

$$ET_i = \frac{1}{i\mu}.$$

Now the time to extinction,  $T$ , of a population of initial size  $N_0$ , is the sum of times between the  $N_0$  successive deaths. i.e.,  $T = T_{N_0} + T_{N_0-1} + \dots + T_1$ . Thus

$$ET = \sum_{i=1}^{N_0} ET_i = \frac{1}{\mu} \sum_{i=1}^{N_0} \frac{1}{i}.$$

The fact that these two results must be equal establishes the identity

$$\sum_{i=1}^n \frac{1}{i} = \sum_{i=1}^n (-1)^i \binom{n}{i} \frac{1}{i}$$

for every positive integer  $n$ . For other proofs of this identity see Feller (1950), p. 63.

#### REFERENCE

Feller, William (1950). An Introduction to Probability Theory and Its Applications. 2nd ed. John Wiley and Sons, New York.