

**SAS PROC GLM and PROC MIXED
for Recovering Inter-Effect Information**

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9 Abstract

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11 It appears that computer programming to recover various types of design effect information is not
12 well understood. In light of the recent development of computer software for performing needed
13 calculations and new statistical procedures, a description of SAS programs is given for experimenters
14 who require such analyses. In the present paper, we describe programs for recovering interblock,
15 interrow, intercolumn, intergradient, and interregression information from incomplete block and lattice
16 rectangle-designed experiments.

17
18 1. Introduction

19 When analyzing the results from experiments designed as incomplete blocks or as lattice
20 rectangles, the information obtained by recovery of interblock or interrow-intercolumn information should
21 be utilized. Ignoring this type of information is akin to ignoring whole plot information in split plot
22 designs. Further, it has been demonstrated that the expected mean squared error of means is smaller in
23 the recovery analyses than in the standard intrablock or intrarow-intracolumn analyses. Also, the
24 resulting means are less affected by random variation among the incomplete blocks or among rows and
25 columns.

26 For certain types of spatial variation, standard textbook analyses may be inappropriate. An
27 example is the differential gradient method within incomplete blocks (or rows) as described by Federer
28 (1996). Other examples include row and column polynomial regressions for row-column designs and the
29 same regressions applied to incomplete blocks in a rectangular array. These regressions vary from
30 complete block to complete block and are considered to be random effects.

31 This paper shows how to program PROC GLM (SAS Institute Inc., 1989) and PROC MIXED
32 (SAS Institute Inc., 1996) to perform the calculations for recovering the above types of information. Since

1 these calculations are easily accomplished, experimenters can routinely perform them instead of using the
2 less efficient intra-effect analyses.

3 2. Recovering Interblock Information

4 We utilize PROC GLM to obtain an ANOVA and intrablock means, and PROC MIXED to
5 compute means adjusted for both intrablock and interblock information. We present a portion of the
6 programs related to class and model. The first part including data, infile, and input is omitted here.

7 The first step is to construct a SAS data set, which for this example is named BLOCK. SAS data
8 sets are rectangular arrays, with columns corresponding to variables and rows to observations. The data
9 set named BLOCK is assumed to contain the following variables: Y (the response), T (the treatment), R
10 (the complete block or replicate), B (the incomplete block within each complete block).

11 A PROC GLM program to obtain an ANOVA and intrablock treatment means is as follows:

```
12     proc glm data=block;
13         class t r b;
14         model y=t r b(r);
15         random r b(r)
16         lsmeans t;
17     run;
```

18
19 The PROC GLM statement invokes the procedure and the DATA=option specifies the analysis
20 data set to be BLOCK. The CLASS statement lists which variables are to be treated as classification
21 variables (as opposed to quantitative variables). Dummy indicator variables are created for each separate
22 level of the classification variables.

23 The MODEL statement specifies the response variable Y and the fixed effects T, R, and B(R), the
24 last one denoting the nesting of B within R. The RANDOM statement requests that R and B(R) be
25 considered random effects in PROC GLM's construction of expected mean squares, although they are still
26 considered as fixed effects in the ANOVA and in the construction of the adjusted means.

27 The LSMEANS statement requests least-squares means of the treatment effect T, which are the
28 intrablock adjusted means.

1 This analysis extends the previous one by the addition of the column effect C(R). The output is the same
 2 as in the previous analysis, and the LSMEANS statement computes the intrarow-intercolumn adjusted
 3 treatment means.

4 A PROC MIXED program for obtaining treatment means adjusted for interrow and intercolumn
 5 information is obtained as follows:

```
6         proc mixed data=row_col;
7             class t r b c;
8             model y=t;
9             random r b(r) c(r);
10            lsmeans t;
11            run;
```

12 REML is used to compute the estimates for the variance components corresponding to R, B(R),
 13 C(R), and the residual. The LSMEANS statement automatically computes the interrow-intercolumn
 14 adjusted treatment means.

16 4. Recovery of Interblock and Intergradient Information

17 When gradients appear within incomplete blocks or within each row (column) of a lattice
 18 rectangle designed experiment, the above analyses will be inappropriate (Federer, 1996). If the gradients
 19 are considered to be random effects, then intergradient information as well as interblock (row)
 20 information will need to be recovered.

21 The nature of the gradients may take various forms. If they follow a polynomial regression
 22 format, then the polynomial regression effects should be included in the SAS data set along with the other
 23 design variables.

24 For this example we assume a linear form for the gradient. The data set GRADIENT is assumed
 25 to contain the following variables: Y (the response), R (the replicate), B (the block within replicate), and
 26 G (the gradient within B).

27 A PROC GLM program to compute the intrablock and intragradient analysis is as follows:

```
28         proc glm data=gradient;
29             class r t b;
30             model y=r t b(r) g g*r g*b(r);
31             random r b(r);
32             lsmeans t;
33             run;
```

1 Note that the gradient variable G is not a class variables, and it enters the model as a main effect and as
 2 an interaction with replications and blocks. (The symbol * is used to indicate interaction.) The particular
 3 model denoted by the above equation assumes that there is a single linear gradient (G) for the entire
 4 experiment, that there are differential linear gradients (G*R) within *each* complete block, and that there
 5 are differential regressions (G*B(R)) within *each* incomplete block. For model one of Federer (1996),
 6 omit the terms “g” and “g*r” from the MODEL statement, for the second model omit “g”, and for the
 7 third model omit “g*r”. The statement in MODEL above is not one of the models considered by Federer
 8 (1996). This presentation demonstrates flexibility in obtaining an appropriate analysis for an experiment.
 9 Although G*R and G*B(R) are considered to be random effects; they were not placed in the RANDOM
 10 statement because PROC GLM allows only classification effects there.

11 A PROC MIXED program for obtaining treatment means adjusted for interblock and
 12 intergradient information and using REML solutions for the variance components is obtained as follows:

```
13         proc mixed data=gradient;
14             class r t b;
15             model y=t g;
16             random r b(r) g*r g*b(r);
17             lsmeans t;
18         run;
```

19
 20 Note that G*R and G*B(R) are now considered to be random effects and G a fixed effect.

21 5. Recovery of Interregression Information

22 If there are differential gradients among the rows and among the columns within complete blocks
 23 and interactions of these regressions, it will be necessary to obtain different analyses than those described
 24 previously. This analysis is an alternative to the standard textbook analyses for incomplete block and
 25 lattice rectangle designed experiments.

26 Suppose that linear and quadratic regressions for rows and columns, along with their
 27 interactions, are considered to be appropriate for controlling the particular type of spatial variation present
 28 in the experiment. Then linear and quadratic effects BL and BQ (for rows) and CL and CQ (for columns)
 29 should be present in the input SAS data set, along with same variables used in previous examples. The
 30 data set for this example is called REGRESS.

1 Searle *et al.* (1992) state that REML and ANOVA solutions for variance components are equal
 2 for all balanced sets of data. By “balanced” they mean orthogonal. For the triple lattice example X1.3 of
 3 Federer (1955), which is a partially balanced incomplete block design, REML and ANOVA solutions are
 4 identical. For the balanced lattice square example of Table 12.5 in Cochran and Cox (1957), REML and
 5 ANOVA solutions are different.

6 Federer (1996) shows that quite striking results are sometimes possible when using other-than-
 7 textbook analyses of responses from experiments. Using the analysis in Section 5 on the data from the
 8 balanced lattice square designed experiment of Table 12.5 in Cochran and Cox (1957), the residual error
 9 mean square was essentially halved. Using the analysis of Section 4 on these data resulted in a decrease
 10 of 16% in the residual error mean square, or the equivalent of an additional replicate. This demonstrates
 11 that models appropriate to the experimental situation should be used rather than using standard textbook
 12 analyses all the time.

13 We conclude with a few comments regarding the use of PROC MIXED. First, when all effects in
 14 the RANDOM statement of PROC MIXED share one or more effects in common, it is often more efficient
 15 computationally to “factor out” this common effect into the optional SUBJECT=effect. For example, the
 16 RANDOM statement from Section 4

17 random r b(r) g*r g*b(r);

18 can also be written as

19 random int b g g*b/subject=r;

20 The gain in computational efficiency grows with the number of levels of R.

21 When performing inference, it is often useful to construct single degree of freedom contrasts
 22 among the adjusted means. Simple differences of the means are automatically generated using the DIFF
 23 option in the LSMEANS statement of PROC MIXED, and multiple comparison adjustments accounting
 24 for simultaneous inference are available as well (refer to SAS Institute, Inc. 1996). Custom contrasts can
 25 be constructed with CONTRAST and ESTIMATE statements as demonstrated in Federer (1995) and in
 26 SAS/STAT Inst. Inc. (1959).

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