A MANUAL OF TRANSFORMATIONS (AND GRAPHS)
USEFUL IN STATISTICS
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## ABSTRACT

A manual is given for the two transformations $\log Y$ and $e^{a Y}$. The manual in its final form would provide easy reference to many transformations that can be useful in the statistical analysis of data.

## INTRODUCTION

On many occasions the statistical analysis of data is improved by making a transformation on the data, e.g., $Y \rightarrow \log Y$. The object of this manual is to provide easy reference to information that helps one decide what transformation(s) might be suitable for experimental data. One use of the manual is that when data can be plotted the resulting plot should be compared with the graphs in Part l. From this comparison (including consideration of properties of the curves in Part 1 with the known properties of your data, e.g. limits, maxima, minima, slope, differential equations and some specific values) select a function $y=f(x)$ which appears to fit your data. Then use the transformation(s) suggested for the function chosen and use Part 2 for the properties of that transformation(s).

## Part 1

Fitting Data

Graph of $y=e^{a x}$


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Equation: \(y=e^{a x}(a>0)\)
    \(x=0, y=1\) for all a
        \(y=0, x\) does not exist
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Limits: $\quad y \rightarrow 0$ as $x \rightarrow-\infty$ $\mathrm{y} \rightarrow \infty$ as $\mathrm{x} \rightarrow \infty$

Differential Equations: $\frac{d y}{d x}=a y$ and in general $\frac{d^{n} y}{d x^{n}}=a^{n} y$ ( $n$ positive integer)
Maxima, Minima, Points of Inflexion: There exists no maximum or point of inflexion. The minimum is $x=0$. The function is monotone increasing.

Transformation to make regression linear: For the model $y=e^{a x} \epsilon$, where $\epsilon$ is a multiplicative error term, the transformation $z=l o g Y$ has the model $z=a x+\epsilon^{\prime}$ with $\epsilon^{\prime}=\log \epsilon$. (See page 8 for properties of the transformation $z=\log Y$.

Transformation to get homogeneous variance: When $Y=e^{a X}$ with $X \sim\left(\mu_{X}, \sigma_{X}^{2}\right)$, $Y \sim\left[e^{a \mu_{X}}, a^{2} \sigma_{x}^{2}\left(e^{a \mu_{X}}\right)^{2}\right]$ with the variance dependent on the mean. But $\log Y \sim\left(a \mu_{X}, a^{2} \sigma_{X}^{2}\right)$ (See page 8 for properties of the transformation log Y.)

Values of $y$ for certain $x$ and $a$

$$
y=e^{a x}
$$

|  | $a$ | 0.5 | 1 | 2 | 4 | 10 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| -10 | 0.368 | 0.007 | 0.00005 | $\rightarrow 0$ | $\rightarrow 0$ | $\rightarrow 0$ |
| -8 | 0.449 | 0.018 | 0.0003 | $\rightarrow 0$ | $\rightarrow 0$ | $\rightarrow 0$ |
| -6 | 0.549 | 0.0498 | 0.0025 | $\rightarrow 0$ | $\rightarrow 0$ | $\rightarrow 0$ |
| -4 | 0.670 | 0.135 | 0.018 | 0.0003 | $\rightarrow 0$ | $\rightarrow 0$ |
| -2 | 0.819 | 0.368 | 0.135 | 0.018 | 0.0003 | $\rightarrow 0$ |
| 0 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 |
| 1 | 1.105 | 1.649 | 2.718 | 7.389 | 54.60 | 22026 |
| 2 | 1.221 | 2.718 | 7.389 | 54.60 | 2981.0 | $\rightarrow \infty$ |
| 4 | 1.492 | 7.389 | 54.60 | 2981.0 | $\rightarrow \infty$ | $\rightarrow \infty$ |
| 6 | 1.822 | 20.086 | 403.43 | $\rightarrow \infty$ | $\rightarrow \infty$ | $\rightarrow \infty$ |
| 8 | 2.2255 | 54.60 | 2981.0 | $\rightarrow \infty$ | $\rightarrow \infty$ | $\rightarrow \infty$ |
| 10 | 2.718 | 148.41 | 22026 | $\rightarrow \infty$ | $\rightarrow \infty$ | $\rightarrow \infty$ |

$$
\text { Graph of } y=\frac{\log x}{a}
$$



Equation: $y=\frac{\log x}{a}(a>0)$

$$
\begin{aligned}
& x=0, y \text { does not exist } \\
& y=0, x=1 \text { for all a }
\end{aligned}
$$

Limits: $\quad y \rightarrow-\infty$ as $x \rightarrow 0$
$\mathrm{y} \rightarrow \infty$ as $\mathrm{x} \rightarrow \infty$
Differential Equations: $\frac{d y}{d x}=\frac{1}{a x}$ and $\frac{d^{n} y}{d x^{n}}+\frac{1}{x} \frac{d^{n-1} y}{d x^{n-1}}=0$ for all integers $n>1$.
Maxima, Minima, Points of Inflexion: There exist none for $x>0$.
Continuity: $y=\frac{\log x}{a}$ is continuous for all positive $x$.
Maximum Gradient: As $x$ approaches 0 the slope of $\frac{\log x}{a}=y$ increases without limit.
Transformation to make regression linear: For the model $y=\frac{\log x}{2}+\epsilon$ the transformation $z=e^{a Y}$ has the model $z=x \epsilon^{\prime}$. with $\epsilon^{\prime}=e^{a \epsilon}$, where $\epsilon$ and $\epsilon^{\prime}$ are error terms. (See page 9 for properties of the transformation $z=e^{a Y}$.

Transformation to get homogeneous variance: When $Y=\frac{\log X}{a}$ with $X \sim\left(\mu_{X}, \sigma_{X}^{2}\right)$ $Y \sim\left(\frac{\log \mu_{x}}{a}, \frac{\sigma_{X}^{2}}{a^{2} \mu_{x}^{2}}\right)$ with the variance dependent on the mean. But $e^{a Y} \sim\left(\mu_{x}, \sigma_{x}^{2}\right)$. (See page 9 for properties of the transformation $e^{a Y}$.)

Values of $y$ for certain $x$ and $a$

$$
y=\frac{\log x}{a}:
$$

| $x$ | 0.1 | 0.5 | 1 | 2 | 4 | 10 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | -- | -- | -- | -- | -- | -- |
| 0.1 | -23.03 | -4.61 | $-2.30$ | -1.15 | -0.58 | -0.23 |
| 0.5 | -6.93 | -1.386 | -0.693 | -0.347 | -0.173 | -0.069 |
| 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2 | 6.93 | 1.386 | 0.693 | 0.347 | 0.173 | 0.069 |
| 4 | 13.86 | 2.77 | 1.386 | 0.693 | 0.347 | 0.139 |
| 6 | 17.92 | 3.584 | 1.792 | 0.8959 | 0.448 | 0.179 |
| 8 | 20.79 | 4.159 | 2.079 | 1.0397 | 0.520 | 0.208 |
| 10 | 23.026 | 4.6052 | 2.303 | 1.151 | 0.576 | 0.230 |
| 20 | 29.957 | 5.991 | 2.996 | 1.498 | 0.749 | 0.2996 |
| 30 | 34.012 | 6.802 | 3.401 | 1.701 | 0.850 | 0.340 |
| 40 | 36.889 | $7 \cdot 378$ | 3.689 | 1.84 | 0.922 | 0.369 |

## Part 2

## Transformations

## Transformation: Log Y <br> mannornanurn

Moments of $\log Y$, given the moments of $Y$ :
For

$$
\begin{gathered}
Y \sim\left(\mu_{Y}, \sigma_{Y}^{2}\right) \\
\log Y \dot{\sim}\left(\log \mu_{Y}, \frac{\sigma_{Y}^{2}}{\mu_{Y}^{2}}\right)
\end{gathered}
$$

using l-term Taylor expansion.
Or

$$
E(\log Y) \dot{\sim} \log \mu_{Y}-\frac{\sigma_{Y}^{2}}{2 \mu_{Y}^{2}}
$$

using 2-term Taylor expansion.

## Distribution of $\log Y$, given distribution of $Y$ :

Let $F(y)$ be the distribution function of $Y$.
Then $G(y)=F\left(e^{y}\right)$ is the distribution function of $\log Y$.

Transformation: $e^{a Y} \quad a>0$
Moments of $e^{a Y}$, given the moments of $Y$ :
For

$$
\begin{gathered}
Y \sim\left(\mu_{Y}, \sigma_{Y}^{2}\right) \\
e^{a Y} \sim\left(e^{a \mu_{Y}}, a^{2} e^{2 a \mu_{Y}} \sigma_{Y}^{2}\right)
\end{gathered}
$$

using l-term Taylor expansion.
Or

$$
E\left(e^{a Y}\right) \dot{\sim} e^{a \mu_{Y}}\left(1+\frac{a^{2}}{2} \sigma_{Y}^{2}\right)
$$

using 2-term Taylor expansion.
Distribution of $e^{a Y}$, given distribution of $Y$ :
Let $F(y)$ be the distribution function of $Y$.
Then $G(y)=F\left(\frac{\log Y}{a}\right)$ is the distribution function of $e^{a Y}$.
(i) When $Y$ has the exponential distribution with parameter $\lambda$,

$$
F(y)= \begin{cases}1-e^{-\lambda y}, & y \geq 0 \\ 0, & y<0\end{cases}
$$

$e^{a Y}$ has the distribution

$$
G(y)=F\left(\frac{\log y}{a}\right)= \begin{cases}1-y^{-\lambda / a}, & y \geq 1 \\ 0, & y<1\end{cases}
$$

with

$$
\begin{gathered}
E\left(e^{a Y}\right)= \begin{cases}\frac{\lambda}{\lambda-a}, & \lambda>a \\
\infty, & \lambda \leq a\end{cases} \\
\operatorname{Var}\left(e^{a Y}\right)= \begin{cases}\frac{\lambda a^{2}}{\left(\lambda-2 a \cdot(\lambda-a)^{2}\right.}, & \lambda>2 a \\
\infty & \lambda \leq 2 a\end{cases}
\end{gathered}
$$

(ii) When $Y$ has the normal distribution with mean. $\mu_{Y}$ and variance $\sigma_{Y}^{2}$ $e^{a Y}$ has the lognormal distribution

$$
G(y)= \begin{cases}\frac{1}{a \sigma_{Y} y \sqrt{2 \pi}} e^{-\frac{\left(\log y-a \mu_{Y}\right)^{2}}{2 a^{2} \sigma_{Y}^{2}}}, & y>0 \\ 0 & y \leq 0\end{cases}
$$

with

$$
\begin{aligned}
E\left(e^{a Y}\right) & =e^{\left(a \mu_{Y}+\frac{1}{2} a^{2} \sigma_{Y}^{2}\right)} \\
\operatorname{Var}\left(e^{a Y}\right) & =e^{\left(2 a \mu_{Y}+2 a^{2} \sigma_{Y}^{2}\right)} e^{\left(2 a \mu_{Y}+a^{2} \sigma_{Y}^{2}\right)}
\end{aligned}
$$

(iii) When $Y$ has the uniform distribution on the interval ( $y_{1}, y_{2}$ ) $e^{a Y}$ has the distribution

$$
G(y)= \begin{cases}\frac{\log y-a y_{1}}{a\left(y_{2}-y_{1}\right)}, & e^{a y_{1}} \leq y \leq e^{a y_{2}} \\ 0 & \text { elsewhere }\end{cases}
$$

with

$$
\begin{aligned}
E\left(e^{a Y}\right) & =\frac{e^{a y_{2}}-e^{a y_{1}}}{a\left(y_{2}-y_{1}\right)} \\
\operatorname{Var}\left(e^{a Y}\right) & =\frac{e^{2 a y_{2}}-e^{2 a y_{1}}}{2 a\left(y_{2}-y_{1}\right)}-\frac{1}{a^{2}\left(y_{2}-y_{1}\right)^{2}}\left(e^{a y_{2}}-e^{a y_{1}}\right)^{2}
\end{aligned}
$$

