BU-80-M

D. S. Robson

In the paper entitled "An empirical law describing heterogeneity in the yields of agricultural crops", H. F. Smith has a result on page 8 which is not immediately obvious. The result referred to is:

$$\left\{ \delta(\log_{e} s^{2}) \right\}^{2} = \left\{ \frac{2s\delta s}{s^{2}} \right\}^{2} = \frac{2}{n}$$

where $\delta s = s/\sqrt{2n} \dots$ ".

This result appears to be obtained by application of the following method for approximating the variance of a function f(X) of a chance variable X. Let μ_X be the mean value of the chance variable X and denote the deviation of X from its mean value by ϵ_X , so that $X = \mu_X + \epsilon_X$. Similarly, let μ_f be the mean value of the function f(X) and $\epsilon_{f(X)} = f(X) - \mu_f$. Now the Taylor series expansion of $f(X) = f(\mu_X + \epsilon_X)$ about the point μ_X is

$$f(\mu_{x}+\epsilon_{x}) = f(\mu_{x}) + f^{\dagger}(\mu_{x}) \epsilon_{x} + f^{\dagger \dagger}(\mu_{x}) \frac{\epsilon_{x}^{2}}{2!} + f^{\dagger \dagger \dagger}(\mu_{x}) \frac{\epsilon_{x}^{3}}{3!} + \dots$$

where $f^{\upsilon}(\mu_x)$ is the v'th derivative of f at the point μ_x . The error $\epsilon_{f(x)}$ therefore has the expansion

(1)
$$\epsilon_{\mathbf{f}(\mathbf{X})} = \mathbf{f}(\mu_{\mathbf{X}} + \epsilon_{\mathbf{X}}) - \mu_{\mathbf{f}} = [\mathbf{f}(\mu_{\mathbf{X}}) - \mu_{\mathbf{f}}] + \mathbf{f}^{\dagger}(\mu_{\mathbf{X}}) \epsilon_{\mathbf{X}} + \mathbf{f}^{\dagger\dagger}(\mu_{\mathbf{X}}) \frac{\epsilon_{\mathbf{X}}}{2^{\dagger}} + \cdots$$

The approximation then consists of dropping all terms on the right side except f'(μ_x) ϵ_y , giving

(2) $\epsilon_{\mathbf{f}(\mathbf{X})} \sim \mathbf{f}^{\dagger}(\mu_{\mathbf{X}}) \epsilon_{\mathbf{X}},$

which is exact only when f(X) is a linear function of X. This gives as an approximation for the variance

(3)
$$\sigma_{\mathbf{f}}^2 = \mathbb{E}\epsilon_{\mathbf{f}(X)}^2 \sim [\mathbf{f}'(\mu_x)]^2 \mathbb{E}\epsilon_X^2 = [\mathbf{f}'(\mu_x)]^2 \sigma_X^2$$
.

In the above example the chance variable X is a sample variance,

$$X = s^{2} = \frac{\sum_{i=1}^{n} (Y_{i} - \overline{y})^{2}}{n-1}$$

and $f(s^2)$ is the natural logarithm,

$$f(X) = f(s^2) = \log_e s^2$$
.

Assuming Y_1,\ldots,Y_n are normal, independent, and identially distributed with variance $\sigma_v^{\ 2}$ we get

$$\mu_{\rm X} = {\rm Es}^2 = \sigma_{\rm Y}^2$$
$$\epsilon_{\rm X} = {\rm s}^2 - \sigma_{\rm Y}^2$$

$$\sigma_{X}^{2} = E\epsilon_{X}^{2} = \frac{2}{n-1}\sigma_{Y}^{4}$$
$$f'(\mu_{x}) = \frac{d(\log_{e}\sigma_{Y}^{2})}{d(\sigma_{Y}^{2})} = \frac{1}{\sigma_{Y}^{2}}$$

so that

$$\sigma_{f}^{2} \sim \left[f'(\mu_{x})\right]^{2} \sigma_{x}^{2} = \left[\frac{1}{\sigma_{y}^{2}}\right]^{2} \frac{2}{n-1} \sigma_{y}^{4} = \frac{2}{n-1}$$

This is essentially the answer arrived at by Smith, but we cannot be certain that this is the method that he used. His assertation

(4)
$$\delta \log_e \sigma_Y^2 = \frac{2\sigma_Y}{\sigma_Y^2} \delta \sigma_Y$$

suggests that he intends δ to be the operator $d/d\sigma_{Y'}$ but if that's the case then $\delta\sigma_{Y} = 1$, contrary to his second assertion

$$\delta \sigma_{\rm Y} = \sqrt{\frac{\sigma_{\rm Y}}{2n}}$$

Because of the resemblance in (4) to the operation of differentiation we conclude that the Taylor series approximation outlined above was originally employed to give the answer 2/n and that the derivation given in Smith's paper is erroneous.

As indicated earlier, this method for approximating the variance of f(X) is exact only when f is a linear function of X. Thus, when

$$f(X) = aX + b$$

$$\mu_{f} = a\mu_{x} + b = f(\mu_{x})$$

$$f'(\mu_{x}) = a$$

$$f^{\vartheta}(\mu_{x}) = 0 \text{ for } \upsilon > 1$$

so that, from (1)

$$\epsilon_{f(X)} = f(\mu_{x} + \epsilon_{X}) - f_{\mu} = f'(\mu_{x}) \epsilon_{X}$$

giving in place of (3) the equality

$$\sigma_{f}^{2} = E\epsilon_{f(X)}^{2} = [f'(\mu_{x})]^{2} E\epsilon_{X}^{2} = a^{2}\sigma_{x}^{2}.$$

The errors of approximation committed when f(X) is a second degree polynomial in X is easily computed. Let

$$f(X) = aX^2 + bX + c$$

then

$$\mu_{f} = a(\sigma_{x}^{2} + \mu_{x}^{2}) + b\mu_{x} + c$$

$$f(\mu_{x}) - \mu_{f} = a\sigma_{x}^{2}$$

$$f'(\mu_{x}) = 2a\mu_{x} + b$$

$$\hat{f''}(\mu_{x}) = 2a$$

$$\hat{f^{0}(\mu_{x})} = 0 \quad \text{for } \mu > 2.$$

Hence, by (1),

$$\epsilon_{f(X)} = a\sigma_{x}^{2} + (2a\mu_{x}+b)\epsilon_{x} + (2a)\frac{\epsilon_{x}}{2!}$$

so the variance σ_f^2 of f(X) is

$$\sigma_{f}^{2} = a^{2}\sigma_{x}^{4} + (2a\mu_{x}+b)^{2}\sigma_{x}^{2} + a^{2}E\epsilon_{x}^{4} + 2a^{2}\sigma_{x}^{4} + 2a(2a\mu_{x}+b)E\epsilon_{x}^{3}$$

₂ 2

while the approximation (3) gives

$$\sigma_{\rm f}^{\ 2} \sim (2 a \mu_{\rm x}^{\ +b})^2 \sigma_{\rm x}^{\ 2}$$
 .

Thus, if X is normally distributed, the approximation underestimates the true variance by an amount $6a^2\sigma_x^{4}$. Clearly, this method of approximation must be used with caution; it is always necessary to verify that the terms being ignored are truly negligible. When X is the maximum likelihood estimator of μ_x , based upon a sample of size n, then the above approximation may be expected to improve as n increases.