BU-224-M<br>U. B. Paik and W. T. Federer<br>July, 1966 Cornell University

ABSTRACT

A method of constructing fractional replicates from a complete factorial is developed and illustrated in the present paper. Special reference is made to the construction of saturated fractional replicates for a specified set of parameters. The method of construction involves a special ordering of the treatment observations and of the single degree of freedom parameter contrasts. Prior to presentation of the method, a generalized inverse method is used to obtain some previous results of Banerjee and Federer on the estimates of parameters and the corresponding variances. Also, a Kronecker product representation is given for the design matrix of any n-factor factorial composed of linear contrasts. Various saturated main effect plans for a $2^{4}$ and a $3^{3}$ factorial are presented. The efficiency of saturated main effect plans is investigated with actual plans being presented for $n=3,4$, and 5 in the $2^{n}$ factorial.

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## SUMMARY

A method of constructing fractional replicates from a complete factorial is developed and illustrated in the present paper. Special reference is made to the construction of saturated fractional replicates for a specified set of parameters. The method of construction involves a special ordering of the treatment observations and of the single degree of freedom parameter contrasts. Prior to presentation of the method, a generalized inverse method is used to obtain some previous results of Banerjee and Federer on the estimates of parameters and the corresponding variances. Also, a Kronecker product representation is given for the design matrix of any n-factor factorial composed of linear contrasts. Various saturated main effect plans for a $2^{4}$ and a $3^{3}$ factorial are presented. The efficiency of saturated main effect plans is investigated with actual plans being presented for $n=3,4$, and 5 in the $2^{n}$ factorial.

## 1. INIRODUCTIION

Raktoe [1966] has shown how to obtain unsaturated and saturated non-orthogonal main effect and resolution $V$ plans---using a single replicate of a lattice design for $2^{n}$ treatments in incomplete blocks of size two. A special ordering of the $2^{\text {n-1 }}$ incomplete blocks was used. Then, from this ordering he obtained the above

[^0]listed fractional replicaces. It is the purpose of this paper to present a method of construction of saturated and unsaturated fractional replicates for any specified set of parameters.

First we shall need to develop and define a notation. Then, some of the results of Banerjee and Federer [1963, 1964, 1966] on the estimates of parameters and their variances will be obtained using a generalized inverse procedure. This alternative development may be useful in other connections. In the next section the single degree of freedom contrast design matrix will be presented as a Kronecker product of the linear contrasts of the levels of each of the $n$ factors. Special orderings of the observations and of the parameter contrasts are used in this Kronecker representation. With this notation developed, the method of construction of fractional replicates is then developed and illustrated with several examples. Various saturated non-orthogonal main effect plans for a $2^{4}$ and a $3^{3}$ factorial are given. Lastly, the efficiency of saturated fractional replicates is investigated; the most efficient (in the sense discussed by Banerjee and Federer [1963, 1966]) non-orthogonal saturated main effect plans are given for $n=3$, 4, and 5 in the $2^{n}$ factorial system.
2. NOTATION

Let $Y$ represent a column vector of $N$ random variables $y_{1},{ }_{2}{ }_{2}, \ldots, y_{N}$, let $B$ represent a column vector of $N$ unknown parameters $b_{I}, b_{2}, \cdots, b_{N}$, and let the known linear orthogonal comparison matrix $X$ (treatment design matrix) in a factorial be composed of $N$ rows and $N$ columns. Then, the observational equation may be represented as:

$$
\begin{equation*}
Y=X B+e, \tag{2.1}
\end{equation*}
$$

where $e$ is an $N \times I$ column vector of random error components, $e_{1}, e_{2}, \cdots, e_{N}$,
where $E(Y)=X B$, and where $E\left(e e^{\prime}\right)=I \sigma^{2}$.
Let $\quad X=\left[\begin{array}{l}X_{1} \\ X_{2}\end{array}\right]$
where $X_{1}$ is a $p \times N$ matrix and $X_{2}$ is an ( $N-p$ ) $\times N$ matrix.
Consider the following fraction

$$
\begin{equation*}
Y_{p}=X_{1} B+e_{p} \tag{2.2}
\end{equation*}
$$

and let $X_{1}=\left[\begin{array}{ll}X_{11} & X_{12}\end{array}\right]$. $B=\left[\begin{array}{l}B_{p} \\ B_{N-p}\end{array}\right]$, and $X_{2}=\left[\begin{array}{lll}X_{21} & X_{22}\end{array}\right]$, where $Y_{p}$ and $B_{p}$ are
$p \times I$ column vectors, $X_{11}$ is a nonsingular $p \times p$ matrix, $X_{12}$ and $X_{21}^{\prime}$ are $p \times(N-p)$, $B_{N-p}$ is an ( $N-p$ ) $\times 1$ column vector, and $X_{22}$ is an ( $\left.N-p\right) \times(N-p)$ matrix, then

$$
Y_{p}=\left[\begin{array}{ll}
X_{11} & X_{12}
\end{array}\right]\left[\begin{array}{l}
B_{p}  \tag{2,3}\\
B_{N-p}
\end{array}\right]+\therefore e_{p}
$$

## 3. USE OF GENERALIZED INVERSE

Using the least squares method, the matrix expression of the normal equations for the fractional replicate given by equation (2.3) is:

$$
\begin{align*}
& {\left[\begin{array}{ll}
X_{11} & X_{12}
\end{array}\right]^{\prime}\left[\begin{array}{ll}
X_{11} & X_{12}
\end{array}\right]\left[\begin{array}{l}
\hat{B}_{p} \\
\hat{B}_{N-p}
\end{array}\right]=\left[\begin{array}{ll}
x_{11} & X_{12}
\end{array}\right]^{\prime} Y_{p}} \\
& {\left[\begin{array}{ll}
X_{11}^{\prime} X_{11} & X_{11}^{\prime} X_{12} \\
X_{12}^{\prime} X_{11} & X_{12}^{\prime} X_{12}
\end{array}\right] \quad\left[\begin{array}{l}
\hat{B}_{p} \\
\hat{B}_{N}-p
\end{array}\right]=\left[\begin{array}{l}
X_{11}^{\prime} \\
X_{12}^{\prime}
\end{array}\right] Y_{p}} \tag{3.1}
\end{align*}
$$

One of the generalized inverses $G$ of $\left[\begin{array}{ll}X_{11}^{\prime} X_{11} & X_{11}^{\prime} X_{12} \\ X_{12}^{\prime} X_{11} & X_{12}^{\prime} X_{12}\end{array}\right]$ is

$$
G=\left[\begin{array}{cc}
\left(X_{11}^{\prime} X_{11}\right)^{-1} & 0  \tag{3.2}\\
0 & 0
\end{array}\right]
$$

for which $H=G X_{1} X_{1}=\left[\begin{array}{cc}I & \left(X_{11}^{\prime} X_{11}\right)^{-1} X_{11}^{\prime} X_{12} \\ 0 & 0\end{array}\right]$

The proof of (3.2) follows easily, i.e.,
$\left[\begin{array}{ll}\mathrm{X}_{11}^{\prime} \mathrm{X}_{11} & \mathrm{X}_{11}^{\prime} \mathrm{X}_{12} \\ \mathrm{X}_{12}^{\prime} \mathrm{X}_{11} & \mathrm{X}_{12}^{\prime} \mathrm{X}_{12}\end{array}\right] G\left[\begin{array}{cc}\mathrm{X}_{11}^{\prime} \mathrm{X}_{11} & \mathrm{X}_{11}^{\prime} \mathrm{X}_{12} \\ \mathrm{X}_{12}^{\prime} \mathrm{X}_{11} & \mathrm{X}_{12}^{\prime} \mathrm{X}_{12}\end{array}\right]=\left[\begin{array}{cc}\mathrm{X}_{11}^{\prime} \mathrm{X}_{11} & \mathrm{X}_{11}^{\prime} \mathrm{X}_{12} \\ \mathrm{X}_{12}^{\prime} \mathrm{X}_{11}\left(\mathrm{X}_{11}^{\prime} \mathrm{X}_{11}\right)^{-1_{X_{11}^{\prime}}^{\prime} \mathrm{X}_{11}} \mathrm{X}_{12}^{\prime} \mathrm{X}_{11}\left(\mathrm{X}_{11}^{\prime} \mathrm{X}_{11}\right)^{-1} \mathrm{X}_{11}^{\prime} \mathrm{X}_{12}\end{array}\right]$
$X_{12}^{\prime} X_{11}\left(X_{11}^{\prime} X_{11}\right)^{-1} X_{11}^{\prime} X_{11}=X_{12}^{\prime} X_{11}$, and since $X_{11}$ is nonsingular, then

$$
\mathrm{X}_{11}\left(\mathrm{X}_{11}^{\prime} \mathrm{X}_{11}\right)^{-1}=\mathrm{X}_{11}^{\prime}{ }^{-1_{X_{11}^{\prime}}^{\prime} \mathrm{X}_{11}\left(\mathrm{X}_{11}^{\prime} \mathrm{X}_{11}\right)^{-1}=\mathrm{X}_{11}^{\prime}-1}
$$

then

$$
\mathrm{X}_{12}^{\prime} \mathrm{X}_{11}\left(\mathrm{X}_{11}^{\prime} \mathrm{X}_{11}\right)^{-1} \mathrm{X}_{12} \mathrm{X}_{12}=\mathrm{X}_{12}^{\prime} \mathrm{X}_{12}
$$

Hence $G$ is a generalized inverse of $\left[\begin{array}{ll}\mathrm{X}_{11}^{\prime} \mathrm{X}_{11} & \mathrm{X}_{11}^{\prime} \mathrm{X}_{12} \\ \mathrm{X}_{12}^{\prime} \mathrm{X}_{11} & \mathrm{X}_{12}^{\prime} \mathrm{X}_{12}\end{array}\right]$.
Therefore

$$
\left[\begin{array}{l}
\hat{B}_{p}  \tag{3.4}\\
\hat{B}_{N-p}
\end{array}\right]=G\left[\begin{array}{l}
X_{11}^{\prime} \\
X_{12}^{\prime}
\end{array}\right] Y_{p}+\left(H-I_{N \times N}\right) Z
$$

$$
=\left[\begin{array}{c}
\left(X_{11}^{\prime} X_{11}\right)^{-I} X_{11}^{\prime}  \tag{3.5}\\
0
\end{array}\right] Y_{p}+\left[\begin{array}{c}
\left(X_{11}^{\prime} X_{11}\right)^{-I} X_{11}^{\prime} X_{12} \\
-I{ }_{(N-p)}(N-p)
\end{array}\right] Z
$$

where $Z$ is an ( $N-p$ ) $\times I$ column vector of arbitrary components $z_{1}, z_{2}, \cdots, z_{N-p}$. Since $Z$ is an arbitrary vector, we can then comsider

$$
\mathrm{Z}=-\hat{\mathrm{B}}_{\mathrm{N} \omega \mathrm{p}} ;
$$

then

$$
\begin{equation*}
\hat{B}_{p}+\left(X_{11}^{\prime} X_{11}\right)^{-1} X_{11}^{\prime} X_{12} \hat{B}_{N-p}=\left(X_{11}^{\prime} X_{11}\right)^{-1} X_{11}^{\prime} Y_{p} \tag{3.6}
\end{equation*}
$$

or

$$
\begin{equation*}
\hat{B}_{p}+X_{11}^{-I} X_{12} \hat{B}_{N-p}=X_{11}^{-I_{Y}} \tag{3.7}
\end{equation*}
$$

Since $\mathrm{X}^{\prime} \mathrm{X}$ is a diagonal matrix, if $\mathrm{X}_{22}{ }^{-1}$ exists, then $X_{11}{ }^{-1}$ exists and we can write X as follows:

$$
\mathrm{X}=\left[\begin{array}{ll}
\mathrm{x}_{11} & \mathrm{X}_{12} \\
\lambda^{\prime} \mathrm{x}_{11} & \mathrm{x}_{22}
\end{array}\right] \text {, where } \lambda=-\mathrm{X}_{12} \mathrm{X}_{22}^{-1}
$$

$$
\text { Let } \quad U_{1}=\left[\begin{array}{l}
x_{11} \\
\lambda^{\prime} x_{11}
\end{array}\right]
$$

then

$$
\left(U_{1}^{\prime} U_{1}\right)^{-I_{U}}{ }_{I}^{\prime} U_{I}=I_{N X N}
$$

then

$$
\begin{align*}
& \left(U_{1}^{\prime} U_{1}\right)^{-1}\left(X_{11}^{\prime} X_{11}^{\prime} \lambda\right)\left[\begin{array}{l}
x_{11} \\
\lambda_{1}^{\prime} X_{11}
\end{array}\right]=I  \tag{3.8}\\
& \left(U_{1}^{\prime} U_{1}\right)^{-1}\left(X_{11}^{\prime} X_{11}+X_{11}^{\prime} \lambda \lambda^{\prime} X_{11}\right)=I \\
& \left(U_{1}^{\prime} U_{1}\right)^{-1} X_{11}^{\prime}\left(I+\lambda \lambda^{\prime}\right) X_{11}=I,
\end{align*}
$$

then

$$
\begin{equation*}
\left(U_{I}^{\prime} U_{1}\right)^{-1} X_{11}^{\prime}\left(I+\lambda \lambda^{\prime}\right)=X_{11}{ }^{-1} . \tag{3.9}
\end{equation*}
$$

Hence, we can rewrite (3.7) as follows:

$$
\begin{align*}
\hat{B}_{p}+\left(U_{1}^{\prime} U_{1}\right) & )^{-X_{11}^{\prime}}\left(I+\lambda \lambda^{\prime}\right) X_{12} \hat{B}_{N-p}  \tag{3.10}\\
& \left.=\left(U_{1}^{\prime} U_{1}\right)\right)^{-1} X_{11}^{\prime}\left(I+\lambda \lambda^{\prime}\right) Y_{p}
\end{align*}
$$

From S. R. Searle [1966], we note

$$
\operatorname{var}\left[\begin{array}{l}
\hat{B}_{p}  \tag{3.11}\\
\hat{B}_{N-p}
\end{array}\right]=G \sigma^{2}
$$

$$
=\left[\begin{array}{cc}
\left(x_{11}^{\prime} x_{11}\right)^{-1} & 0 \\
0 & 0
\end{array}\right] \sigma^{2} ;
$$

then

$$
\begin{equation*}
\operatorname{var}\left(\hat{B}_{p}\right)=\left(X_{11}^{\prime} X_{11}\right)^{-1 \sigma^{2}} . \tag{3.12}
\end{equation*}
$$

These results are equivalent to results of Banerjee and Federer [1963, 1964].
4. KRONECKER PRODUCT, CONSIRUCTION OF THE DESIGN MATRIX X

The ordering of the treatments in the vector $Y$ is as follows: Set the first n -l factors at the first level and run through all levels of the $\mathrm{n}^{\text {th }}$ factor consecutively; then set all levels of the first $n-2$ factors at the first level and set the level of the $n-1^{8 t}$ factor at the second level and run through all levels of the $n^{\text {th }}$ factor in consecutive order; continue this process until all levels of the $n-I^{s t}$ factor have been exhausted in consecutive order; then run through levels of the $n-2^{n d}$ factor in the manner described for the $n-1^{s t}$ factor; continue this process for the $n-3^{\text {rd }}$ up to and including the first factor which exhausts all the
combinations in the $n$-factor factorial. Consequently, if we suppose the $i^{t h}$ factor has $n_{i}$ levels, then we can express the $k^{t h}$ treatment as follows:

$$
\begin{equation*}
\left(\alpha_{1}, \alpha_{2}, \cdots, \alpha_{h}, \cdots, \alpha_{q}, \cdots, \alpha_{n}\right) \tag{4.1}
\end{equation*}
$$

where

$$
\begin{aligned}
& \alpha_{j}=\left[\frac{k_{j-1}}{\prod_{i=2}^{n} n_{i}}\right] \text { for } j=1, \cdots, n-1 ; \\
& \alpha_{n}=k_{n-1},
\end{aligned}
$$

where

$$
k_{0}=k, \quad k_{j-2} \equiv k_{j-1}\left(\bmod \prod_{i=j}^{n} n_{i}\right),
$$

and
[x] denotes the largest integer less than $x$ or equal to $x$.

The treatment ordering in the examples given below follow this ordering.

The second step is to set up the linear contrast matrix for the $q$ levels of each of the $n$ factors; we shall denote this contrast matrix as $L_{q h}$ where $q$ refers to the number of levels associated with the $h^{t h}$ factor. The third step in the representation of the design matrix is to take the Kronecker product of the linear contrast matrices, i.e. $\prod_{h=1}^{n} \otimes I_{q h}$. The parameter order is such that the mean and the $n^{\text {th }}$ factor contrasts appear first, then the first contrast of the $n-1^{\text {st }}$ factor and interactions with the $n^{\text {th }}$ factor contrasts appear next, etc. Here, we know that the symbols of the factors are arbitrary. If we change the two factor symbols, each other in the linear contrast matrix, for example the $h^{\text {th }}$ factor symbol and the $q^{t h}$ factor symbol, but keep the factor symbols and their order in the treatment, then, in the vector $Y$ the $k^{t h}$ treatment (4.1) will become as follows:

$$
\begin{equation*}
\left(\alpha_{1}, \alpha_{2}, \cdots, \alpha_{q}, \cdots, \alpha_{n}, \cdots, \alpha_{n}\right) \tag{4.2}
\end{equation*}
$$

The examples below illustrate the procedure.

Example 4.1, $3 \times 2$ factorial
Consider a $3 \times 2$ factorial arrangement of treatments, and suppose factor $A$ is represented at the three levels 0,1 , and 2 and factor $B$ at the two levels 0 and 1; then, we obtain the following coefficients for the 6 orthogonal contrasts among 6 treatments by using the Kronecker product of the two matrices $L_{3 A}$ and $L_{2 B}$ (e.g., see Robson [1959]) where

If we represent the matrix of coefficients given above by $I_{6}$, then

$$
I_{6}=I_{3 A} \otimes I_{2 B}
$$

where $\underset{x}{ }$ refers to the Kronecker product. $I_{6}$ is equivalent to the design matrix X of a complete $3 \times 2$ factorial.

Example 4.2, $2^{2}$ factorial

$$
\begin{aligned}
I_{4} & =I_{2 A} \times I_{2 B} \\
\cdots & M \cdot \\
& =\left[\begin{array}{rrrr}
1 & -1 & -1 & 1 \\
1 & 1 & -1 & -1 \\
1 & -1 & 1 & -1 \\
1 & 1 & 1 & 1
\end{array}\right]
\end{aligned}
$$

where the treatment order is $00,01,10$, and 11 , the factors are $A$ and $B$, and the parameter order is $M, B, A$, and $A B$.

Example 4.3, $2^{3}$ factorial

$$
\begin{aligned}
L_{8} & =I_{2 A} \otimes I_{4}=L_{2 A} \otimes I_{2 B} \otimes I_{2 C} \\
& =\left[\begin{array}{cc}
I_{4} & -I_{4} \\
I_{4} & I_{4}
\end{array}\right],
\end{aligned}
$$

where the treatment order is 000, 001, 010, 011, 100, 101, 110, and 111, the factors are $A, B$, and $C$, and the parameter order is $M, C, B, B C, A, A C, A B$, and ABC.

Example 4.4, $3^{2}$ factorial

$$
\begin{aligned}
L_{9} & =I_{3 A} \times I_{3 B} \\
& =\left[\begin{array}{rrrrrrrrr}
1 & -1 & 1 & -1 & 1 & -1 & 1 & -1 & 1 \\
1 & 0 & -2 & -1 & 0 & 2 & 1 & 0 & -2 \\
1 & 1 & 1 & -1 & -1 & -1 & 1 & 1 & 1 \\
1 & -1 & 1 & 0 & 0 & 0 & -2 & 2 & -2 \\
1 & 0 & -2 & 0 & 0 & 0 & -2 & 0 & 4 \\
1 & 1 & 1 & 0 & 0 & 0 & -2 & -2 & -2 \\
1 & -1 & 1 & 1 & -1 & 1 & 1 & -1 & 1 \\
1 & 0 & -2 & 1 & 0 & -2 & 1 & 0 & -2 \\
1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1
\end{array}\right],
\end{aligned}
$$

where the treatment order is $00,01,02,10,11,12,20,21$, and 22 , the factors are $A$ and $B$, and the parameter order is $M, B_{L}, B_{Q}, A_{I}, A_{L} B_{L}, A_{L} B_{Q}, A_{Q}, A_{Q} B_{L}$, and $A_{Q}{ }^{B}{ }_{Q}$

Example 4.5, $3^{3}$ factorial

$$
\begin{aligned}
I_{27} & =I_{3} \circledast I_{9}=I_{3 A} \circledast I_{3 B} \otimes I_{3 C} \\
& =\left[\begin{array}{ccc}
L_{9} & -I_{9} & I_{9} \\
I_{9} & 0 & -2 I_{9} \\
I_{9} & I_{9} & I_{9}
\end{array}\right]
\end{aligned}
$$

where the treatment order is 000, 001, 002, 010, 011, 012, 020, 021, 022; 100, 101, 102, 110, 111, 112, 120, 121, 122; 200, 201, 202, 210, 211, 212, 220, 221, and 222, the factors are $A, B$, and $C$, and the parameter order is $M, C_{L}, C_{Q}, B_{L}, B_{L} C_{L}, B_{L} C_{Q}$, $B_{Q}, B_{Q} C_{L}, B_{Q} C_{Q} ; A_{L}, A_{L} C_{L}, A_{L} C_{Q}, A_{L} B_{L}, A_{L} B_{L} C_{L}, A_{L} B_{L} C_{Q}, A_{L} B_{Q}, A_{L} B_{Q} C_{L}, A_{L} B_{Q} C_{Q} ; A_{Q}$, $A_{Q} C_{I}, A_{Q} C_{Q}, A_{Q} B_{L}, A_{Q} B_{L} C_{L}, A_{Q} B_{L} C_{Q}, A_{Q} B_{Q}, A_{Q} B_{Q} C_{L}$, and $A_{Q} B_{Q} C_{Q}$.

Here, if we change the two factor symbols $A$ and $C$ with each other in the $L_{27}$, then the parameter order is $M, A_{L}, A_{Q}, B_{L}, A_{L} C_{I}, A_{Q} B_{L}, B_{Q}, A_{L} B_{Q}, A_{Q} B_{Q} ; C_{L}, A_{L} C_{L}$, $A_{Q} C_{I}, B_{L} C_{I}, A_{L} B_{L} C_{I}, A_{Q} B_{L} C_{L}, B_{Q} C_{L}, A_{L} B_{Q} C_{L}, A_{Q} B_{Q} C_{L} ; C_{Q}, A_{L} C_{Q}, A_{Q} C_{Q}, B_{L} C_{Q}, A_{L} B_{L} C_{Q}$, $A_{Q} B_{L} C_{Q}, B_{Q} C_{Q}, A_{L} B_{Q} C_{Q}$, and $A_{Q} B_{Q} C_{Q}$, and treatment order is 000, 100, 200, 010, 110, 210, 020, 120, 220; 001, 101, 201, 011, 111, 211, 021, 121, 221; 002, 102, 202, 012, 112, 212, 022, 122, and 222.

## 5. REARRANGING THE TREATMENT ORDER

If we recall the solution (3.7) or (3.10), we will see that we need the inverse of $X_{11}$ or of $X_{22}$ to obtain the solution. Also we shall see in section 6 that if the size of the fraction is less than $s^{n-1}$ in the $s^{n}$ factorial, then we can use the $s^{n-1} \times s^{n-1}$ orthogonal matrix $X_{11}^{*}$ (in the sense of diagonal) instead of $s^{n} \times s^{n}$ matrix $X$ to get a solution such as (3.7) or (3.10). Also we shall see in this case that the method of construction of a saturated fractional replicate is the problem which selects the smallest number of treatments from the treatments
corresponding to the orthogonal matrix $X_{11}^{*}$. Here we also recall that, in section 4, $L_{s_{n}}$ is already an orthogonal matrix, then we can construct a saturated fractional replicate from the first $s^{n-1}$ treatments in the vector $Y$. But, in this case, the mean effect will be confounded with the main effect A. This is the reason that we must rearrange the treatment order in the vector $Y$ with some higher order defining contrast before constructing a fractional plan; i.e., the mean effect is required to be unconfounded with the main effects.

Now consider rearranging of the treatment order in the vector $Y$ with some defining contrast in $s^{n}$ factorial. If we use the expression (4.1) for the treatments, then the numbers $\alpha_{i}$ take on values from 0 to (sml). The $s^{n}-1$ degrees of freedom among the $s^{n}$ treatment combinations may be partitioned into ( $\left.s^{n}-1\right) /(s-1)$ sets of $s-1$ degrees of freedom. Each set of ( $s-1$ ) degrees of freedom is given by the contrast among the $s$ sets of $s^{n-1}$ treatment combinations specified by the following s equations:

$$
\begin{gather*}
\alpha_{1}+c_{2} \alpha_{2}+\cdots+c_{n} \alpha_{n}=0 \\
\alpha_{1}+c_{2} \alpha_{2}+\cdots+c_{n} \alpha_{n}=1  \tag{5.1}\\
\vdots \\
\alpha_{1}+c_{2} \alpha_{2}+\cdots+c_{n} \alpha_{n}=s-1
\end{gather*}
$$

where the right-hand sides of these equations are elements of the Galois Field GF(s). The $c_{i}^{\prime} s$ are positive integers between 0 and ( $s m$ ), not all equal to zero, and all addition and multiplication is done within the Galois Field GF(s), then the interaction $A B^{c_{2}} \ldots K^{c_{n}}$ corresponds to the equation whose left-hand side is $\alpha_{1}+c_{2} \alpha_{2}+\cdots+c_{n} \alpha_{n}$

Consider a defining contrast, i.e.,

$$
\begin{equation*}
I=A B^{c_{2}} \ldots K^{c_{n}} \tag{5.2}
\end{equation*}
$$

(the factor $A$ is always included for convenience) then the identity relationships will be written as

$$
\begin{align*}
I= & \left(A B^{c} \ldots K^{c}{ }^{n}\right)_{0} \\
I= & \left(A B^{c} \ldots K^{c}{ }^{n}\right)_{I}  \tag{5.3}\\
& \vdots \\
I= & \left.\left(A B{ }^{c}{ }^{c} \ldots K^{c}\right)^{n}\right)_{S-I}
\end{align*}
$$

Let, the set of treatments for fixed $\alpha_{1}$ be $\left\{\alpha_{1}:\left(\alpha_{1}, \alpha_{2}, \cdots, \alpha_{n}\right)\right\}$, then, from (4.3) and (4.5), we can find easily the following relationships:


If $s=p^{m}$ ( $p$ is a prime number), this pattern (5.4) will be different but will obtain similar relationships among the treatment numbers in the below.

If we denote

$$
\begin{equation*}
\left\{k_{i}\right\} \tag{5.5}
\end{equation*}
$$

as the set of the treatment numbers corresponding to $\left(A B^{c} 2 \ldots K^{c}\right)_{i}$ in the set $\left\{0:\left(0, \alpha_{2}, \cdots, \alpha_{n}\right)\right\}$, then the whole set of the treatment numbers corresponding to
$\left(A B^{c_{2}} \ldots K^{c}\right)_{0}$ in the vector $Y$ will be

$$
\begin{equation*}
\left(\left\{k_{0}\right\},\left\{s+k_{s-1}\right\},\left\{2 s+k_{s-2}\right\}, \cdots,\left\{(s-1) s+k_{l}\right\}\right) \tag{5.6}
\end{equation*}
$$

If we use the notation

$$
\begin{equation*}
\left\{\underline{k}_{i}\right\} \tag{5.7}
\end{equation*}
$$

as the corresponding set of the row vectors in the $L_{s^{n}}$-I, then we will easily recognize the following:

$$
\begin{gather*}
\left\{\underline{s+k_{s}-1}\right\}=\left\{\underline{k}_{s-1}\right\} \\
\left\{\underline{s+k_{s}} \mathbf{s - 2}\right\}=\left\{\underline{k}_{s-2}\right\}  \tag{5.8}\\
\vdots \\
\left\{(\underline{s-1})_{s+k_{1}}\right\}=\left\{\underline{k}_{1}\right\} .
\end{gather*}
$$

Hence, if we rearrange the treatment order with defining contrast $I=A B^{c} 2 \ldots{ }^{c}{ }^{c}{ }_{n}$ in the vector $Y$, then the first matrix $L^{*}{ }_{n-1}$ corresponding to the set of $\left\{\left(A B^{c} 2 \ldots K^{c} n_{0}\right\}\right.$ has the same row vectors as $I_{S_{n-1}}$ except ordering, i.e.,

$$
L_{s}^{*}{ }_{n-1} \sim L_{s_{n-1}}
$$

where the notation $\sim$ means that if we rearrange the row vector order properly in the $\mathrm{L}_{\mathrm{s}}^{*}{ }_{\mathrm{n}-1}$, then $\mathrm{I}_{\mathrm{s}}^{*}{ }_{\mathrm{n}-1}$ will be the same as $\mathrm{L}_{\mathrm{s}}{ }_{\mathrm{n}-I^{-}}$

Example 5.1, $3^{3}$ Pactorial
Consider the $\frac{1}{3}$ replicate of the $3^{3}$ factorial with the defining contrast $I=A B C^{2}$, then the treatment combinations which comprise the fractional replicate satisfy one of the following equations:

$$
\begin{array}{ll}
\alpha_{1}+\alpha_{2}+2 \alpha_{3}=0 & (\bmod 3) \\
\alpha_{1}+\alpha_{2}+2 \alpha_{3}=1 & (\bmod 3) \\
\alpha_{1}+\alpha_{2}+2 \alpha_{3}=2 & (\bmod 3)
\end{array}
$$

then the identity relationships are written as

$$
\begin{aligned}
& I=\left(A B C^{2}\right)_{0} \\
& I=\left(A B C^{2}\right)_{1} \\
& I=\left(A B C^{2}\right)_{2}
\end{aligned}
$$

From example (4.4) and (5.7)

$$
\begin{aligned}
& \left\{\underline{k}_{0}\right\}=\left[\begin{array}{rrrrrrrrr}
1 & -1 & 1 & -1 & 1 & -1 & 1 & -1 & 1 \\
1 & 0 & -2 & 0 & 0 & 0 & -2 & 0 & 4 \\
1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1
\end{array}\right] \\
& \left\{\underline{k}_{1}\right\}=\left[\begin{array}{rrrrrrrrr}
1 & 1 & 1 & -1 & -1 & -1 & 1 & 1 & 1 \\
1 & -1 & 1 & 0 & 0 & 0 & -2 & 2 & -2 \\
1 & 0 & -2 & 1 & 0 & -2 & 1 & 0 & -2
\end{array}\right] \\
& \left\{\underline{k}_{2}\right\}=\left[\begin{array}{rrrrrrrrr}
1 & 0 & -2 & -1 & 0 & 2 & 1 & 0 & -2 \\
1 & 1 & 1 & 0 & 0 & 0 & -2 & -2 & -2 \\
1 & -1 & 1 & 1 & -1 & 1 & 1 & -1 & 1
\end{array}\right]
\end{aligned}
$$

then, from (5.8)

$$
\mathrm{I}_{9}^{*}=\left[\begin{array}{l}
\left\{\underline{\underline{k}}_{0}\right\} \\
\left\{\underline{\underline{k}}_{2}\right\} \\
\left\{\underline{\underline{k}}_{1}\right\}
\end{array}\right] .
$$

The treatment order corresponding matrix $L_{9}^{*}$ is, from (5.6) and (5.8), as follows: 000, 011, 022, 101, 112, 120, 202, 210, 221.
6. CONSIRUCTION OF FRACIIONAL REPIICATES

We shall consider mostly the method of construction of the saturated main effect plans in $s^{n}$ factorial. Although we could always construct various saturated
non-orthogonal plans for any given parameter set, the general steps of the construction method may not be too instructive. The following steps, however, will be common for the construction of a fractional replicate for given parameters. The special consideration for each case will be illustrated in the following examples.

Step 1. Given the design matrix and the parameters and observation vectors $X B=Y$ in any fashion and not necẹssarily that of the previous section, we now rearrange the parameter matrix such that the p parameters, $p<\mathbb{N}$, are arranged to have the $p$ parameters of interest first and the $N-p$ parameters not of interest to obtain $B^{\prime}$ rearranged as ( $B_{p}^{* \prime} B_{N-p}^{* \prime}$ ). This also rearranges the columns of $X$ such that

$$
\begin{gather*}
X^{*} B^{*}=Y  \tag{6.1}\\
\left(\begin{array}{cc}
X_{1}^{*} & \left.X_{2}^{*}\right) \\
\mathbb{N X p} & \mathbb{N X}(\mathbb{N}-p)
\end{array}\left[\begin{array}{c}
B_{p}^{*} \\
B_{N}^{*}-p
\end{array}\right]=Y\right. \tag{6.2}
\end{gather*}
$$

Step 2. Search through rows of $X_{1}^{*}$ until there is an $X_{11}$, $p \times p$, which is nonsingular.

Step 3. Corresponding to the rows in $X_{11}$ will be rows in $X_{1}^{*}$ and observations in $Y$. Rearrange the observations in $Y$ into

$$
\left[\begin{array}{c}
Y_{p}^{*} \\
Y_{N}^{*}-\mathrm{p}
\end{array}\right]
$$

to correspond to the rows in $X_{11}$ from $X_{1}^{*}$. The observations in $Y_{p}^{*}$ yield a saturated design for the $p$ parameters in $B_{p}^{*}$.

This obtained set is one of the possible sets. All possible sets are found by finding all $X_{11}$ which have an inverse.

Example 6.1, saturated main effect plans in a $2^{4}$ factorial
If we consider a $2^{4}$ factorial design matrix $L_{16}$ with the defining contrast $I=A B C D$, and if we consider only the first eight treatments after rearranging treatment order with defining contrast $I=A B C D$, then the alias scheme is as follows:

$$
\begin{aligned}
& M \doteq A B C D, \quad A \doteq B C D, \quad B \doteq A C D, \quad C \doteq A B D, \quad D \doteq A B C \\
& A B \doteq C D, \quad A C \doteq B D, \quad B C \doteq A D .
\end{aligned}
$$

After rearranging the rows and columns under consideration of the above alias scheme and using property (5.8), we obtain the following matrix $X$ :

$$
X=\left[\begin{array}{ll}
X_{11}^{*} & X_{11}^{*}  \tag{6.3}\\
X_{11}^{*} & -X_{1}^{*}
\end{array}\right]=\left[\begin{array}{ccccccccccccccccc}
1 & -1 & -1 & -1 & -1 & 1 & 1 & 1 & 1 & -1 & -1 & -1 & -1 & 1 & 1 & 1 \\
1 & 1 & 1 & -1 & -1 & -1 & -1 & 1 & 1 & 1 & 1 & -1 & -1 & -1 & -1 & 1 \\
1 & 1 & -1 & -1 & 1 & 1 & -1 & -1 & 1 & 1 & -1 & -1 & 1 & 1 & -1 & -1 \\
1 & -1 & 1 & -1 & 1 & -1 & 1 & -1 & 1 & -1 & 1 & -1 & 1 & -1 & 1 & -1 \\
1 & 1 & -1 & 1 & -1 & -1 & 1 & -1 & 1 & 1 & -1 & 1 & -1 & -1 & 1 & -1 \\
1 & -1 & 1 & 1 & -1 & 1 & -1 & -1 & 1 & -1 & 1 & 1 & -1 & 1 & -1 & -1 \\
1 & -1 & -1 & 1 & 1 & -1 & -1 & 1 & 1 & -1 & -1 & 1 & 1 & -1 & -1 & 1 \\
1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 \\
1 & -1 & -1 & -1 & -1 & 1 & 1 & 1 & -1 & 1 & 1 & 1 & 1 & -1 & -1 & -1 \\
1 & 1 & 1 & -1 & -1 & -1 & -1 & 1 & -1 & -1 & -1 & 1 & 1 & 1 & 1 & -1 \\
1 & 1 & -1 & -1 & 1 & 1 & -1 & -1 & -1 & -1 & 1 & 1 & -1 & -1 & 1 & 1 \\
1 & -1 & 1 & -1 & 1 & -1 & 1 & -1 & -1 & 1 & -1 & 1 & -1 & 1 & -1 & 1 \\
1 & 1 & -1 & 1 & -1 & -1 & 1 & -1 & -1 & -1 & 1 & -1 & 1 & 1 & -1 & 1 \\
1 & -1 & 1 & 1 & -1 & 1 & -1 & -1 & -1 & 1 & -1 & -1 & 1 & -1 & 1 & 1 \\
1 & -1 & -1 & 1 & 1 & -1 & -1 & 1 & -1 & 1 & 1 & -1 & -1 & -1 & 1 & -1 \\
1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & -1 & -1 & -1 & -1 & -1 & -1 & -1 & -1
\end{array}\right]
$$

where the treatment order is
0000, 0011, 0110, 0101, 1010, 1001, 1100, 1111;
1000, 1011, 1110, 1101, 0010, 0001, 0100, and 0111,
and the parameter order is

$$
\begin{align*}
& M, C, D, B C D, B, B C, B D, C D ; \\
& A B C D, A B D, A B C, A, A C D, A D, A C, A B \tag{6.5}
\end{align*}
$$

Consider the following fraction of a $2^{4}$ factorial

$$
\begin{equation*}
Y_{p}=X_{I} B+e_{p}, p<8 \tag{6.6}
\end{equation*}
$$

where $Y_{p}$ is a $p \times I$ vector from the vector $Y, B$ is a column vector of $N=16$ unknown parameters reordered such as (6.5), $X_{1}$ is a design matrix for given $Y_{p}$ and $B$, and $e_{p}$ is a $p \times I$ column vector of random error.

Suppose the following partition matrix of X is possible:

$$
\begin{align*}
\mathrm{x}=\left[\begin{array}{l}
\mathrm{x}_{1} \\
\mathrm{x}_{2}
\end{array}\right] & =\left[\begin{array}{ll}
\mathrm{x}_{11} & \mathrm{x}_{12} \\
\mathrm{x}_{21} & \mathrm{x}_{22}
\end{array}\right] \\
& =\left[\begin{array}{ll:l}
\mathrm{x}_{11} & \mathrm{x}_{1211} & \mathrm{x}_{1212} \\
\mathrm{x}_{2111} & \mathrm{x}_{2211} & x_{2212} \\
\hdashline \mathrm{x}_{2121} & \mathrm{x}_{2221} & x_{2222}
\end{array}\right] \tag{6.7}
\end{align*}
$$

and

$$
X_{11}^{*}=\left[\begin{array}{ll}
x_{11} & x_{1211}  \tag{6.8}\\
x_{2111} & x_{2211}
\end{array}\right]
$$

where $X_{11}$ is a $p \times p(p<8)$ nonsingular matrix, $X_{2111}$ and $X_{1211}^{\prime}$ are each $p \times(8-p)$ matrices, $X_{2111}$ is an ( $8-p$ ) $\times(8-p)$ matrix, $X_{2121}$ and $X_{1212}^{\prime}$ are each $8 \times \mathrm{p}$ matrices, and $\mathrm{X}_{2221}$ and $\mathrm{X}_{2212}^{\prime}$ are each $8 \times(8-\mathrm{p})$ matrices.

We know from (6.3) that

$$
x_{1212}=\left[\begin{array}{l:l}
x_{11} & x_{1211} \tag{6.9}
\end{array}\right]
$$

and since $X_{11}^{*} \mathrm{X}_{11}^{*}$ is diagonal, if $\mathrm{X}_{11}$ is nonsingular, then $\mathrm{X}_{2211}$ is also nonsingular, and from (3.9)

$$
\left.\begin{array}{rl}
\mathrm{X}_{11}^{-1} & =\left(U_{11}^{\prime} U_{11}\right)^{-1} X_{11}^{\prime}\left(1+\lambda_{1} \lambda_{1}^{\prime}\right) \\
\text { where } & U_{11}
\end{array}\right)=\left[\begin{array}{c}
x_{11} \\
x_{2111} \tag{6.12}
\end{array}\right] \quad \text { and } \quad \lambda_{1}=-X_{1211} X_{2211}^{-1} .
$$

then $X_{11}^{\prime}\left(I+\lambda_{1} \lambda_{1}^{\prime}\right) X_{12}=X_{11}^{\prime}\left(I+\lambda_{1} \lambda_{1}^{\prime}\right)\left[X_{1211} \mathrm{X}_{11}: \mathrm{X}_{1211}\right]$

$$
\begin{equation*}
=\left[X_{11}^{\prime}\left(I+\lambda_{1} \lambda_{1}^{\prime}\right) X_{1211} \dot{幺}_{1} 8 I_{5 \times 5} \dot{\vdots}_{11}^{\prime}\left(I+\lambda_{1} \lambda_{1}^{\prime}\right) X_{1211}\right] \tag{6.13}
\end{equation*}
$$

then, from (3.7) and (3.10) we will obtain the following solution for (6.6):

$$
\begin{gather*}
\hat{B}_{p}+\left(U_{11}^{\prime} U_{11}\right)^{-1}\left[X_{11}^{\prime}\left(I+\lambda_{1} \lambda_{1}^{\prime}\right) X_{1211}: 8 I: X_{11}^{\prime}\left(I+\lambda_{1} \lambda_{1}^{\prime}\right) X_{1211}\right] \hat{B}_{16-p} \\
=\left(U_{11}^{\prime} U_{11}\right)^{-1} X_{11}^{\prime}\left(I+\lambda_{1} \lambda_{1}^{\prime}\right) Y_{p} \tag{6.14}
\end{gather*}
$$

This solution indicates clearly that the mean and main effects are not confounded with each other and that the solution depend only on $\lambda_{1}$. This further means that the solution depends only on $X_{2211}$

Now consider the saturated main effect plans in a $2^{4}$ factorial. Let the treatments be arranged such as (6.4) and the corresponding row vectors in X be numbered $1,2, \cdots, 16$ respectively, and let

$$
U_{12}=\left[\begin{array}{l}
X_{1211}  \tag{6.15}\\
X_{2211}
\end{array}\right]=\left[\begin{array}{rrr}
1 & 1 & 1 \\
-1 & -1 & 1 \\
1 & -1 & -1 \\
-1 & 1 & -1 \\
-1 & 1 & -1 \\
1 & -1 & -1 \\
-1 & -1 & 1 \\
1 & 1 & 1
\end{array}\right]
$$

In the matrix $U_{12}$, we can find easily three independent rows, i.e., the following combinations of rows make nonsingular $3 \times 3$ matrices:

$$
\begin{aligned}
& (1,2,3),(1,2,4),(1,2,5),(1,2,6),(1,3,4), \\
& (1,3,5),(1,3,7),(1,4,6),(1,4,7),(1,5,6), \\
& (1,5,7),(2,3,4),(2,3,5),(2,3,6),(2,3,8), \\
& (2,4,8),(2,5,8),(2,6,8),(3,4,7),(3,4,8), \\
& (3,5,7),(3,5,8),(3,7,8),(4,6,7),(4,6,8), \\
& (4,7,8),(5,6,7),(5,6,8),(5,7,8),(6,7,8),
\end{aligned}
$$

where the numbers indicate the row numbers in matrix $U_{12}$, then the following 30 treatment combinations will be the saturated main effect plans in a $2^{4}$ factorial:

| (1) | 0101 | (2) | 0110 | (3) | 0110 | (4) | 0110 | (5) | 0011 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1010 |  | 1010 |  | 0101 |  | 0101 |  | 1010 |
|  | 1001 |  | 1001 |  | 1001 |  | 1010 |  | 1001 |
|  | 1100 |  | 1100 |  | 1100 |  | 1100 |  | 1100 |
|  | 1111 |  | 1111 |  | 1111 |  | 1111 |  | 1111 |
| (6) | 0011 | (7) | 0011 | (8) | 0000 | (9) | 0000 | (10) | 0000 |
|  | 0110 |  | 0110 |  | 1010 |  | 0110 |  | 0110 |
|  | 1010 |  | 0101 |  | 1001 |  | 1010 |  | 0101 |
|  | 1100 |  | 1001 |  | 1100 |  | 1001 |  | 1010 |
|  | 1111 |  | 1110 |  | 1111 |  | 1100 |  | 1100 |
| (11) | 0000 | (12) | 0000 | (13) | 0000 | (14) | 0000 | (15) | 0000 |
|  | 0011 |  | 0011 |  | 0011 |  | 0011 |  | 0011 |
|  | 0101 |  | 0101 |  | 0110 |  | 0110 |  | 0110 |
|  | 1001 |  | 1010 |  | 1010 |  | 1010 |  | 1010 |
|  | 1111 |  | 1001 |  | 1111 |  | 1100 |  | 1001 |
| (16) | 0000 | (17) | 0000 | (18) | 0000 | (19) | 0000 | (20) | 0011 |
|  | 0011 |  | 0011 |  | 0011 |  | 0011 |  | 0101 |
|  | 0110 |  | 0110 |  | 0110 |  | 0110 |  | 1001 |
|  | 0101 |  | 0101 |  | 0101 |  | 0101 |  | 1100 |
|  | 1111 |  | 1100 |  | 1001 |  | 1010 |  | 1111 |
| (21) | 0011 | (22) | 0011 | (23) | 0011 | (24) | 0000 | (25) | 0000 |
|  | 0101 |  | 0110 |  | 0110 |  | 0101 |  | 0101 |
|  | 1010 |  | 1010 |  | 0101 |  | 1001 |  | 1010 |
|  | 1001 |  | 1001 |  | 1100 |  | 1100 |  | 1100 |
|  | 1111 |  | 1111 |  | 1111 |  | 1111 |  | 1111 |
| (26) | 0000 | (27) | 0000 | (28) | 0000 | (29) | 0000 | (30) | 0000 |
|  | 0101 |  | 0110 |  | 0011 |  | 0011 |  | 0011 |
|  | 1010 |  | 0101 |  | 1010 |  | 1010 |  | 0101 |
|  | 1001 |  | 1001 |  | 1001 |  | 1001 |  | 1001 |
|  | 1100 |  | 1100 |  | 1111 |  | 1100 |  | 1100 |
|  | -- |  |  |  |  |  |  |  |  |

Let $\left(n_{1}, n_{2}, n_{3}, n_{4}, n_{5}\right)$ be one of the above 30 plans, then by recalling (5.6), (5.7), and (5.8), we will know the following treatment combinations are also saturated main effect plans in a $2^{4}$ factorial, i.e.,

$$
\begin{equation*}
\left(n_{1}+8, n_{2}+8, n_{3}+8, n_{4}+8, n_{5}+8\right) \tag{6.17}
\end{equation*}
$$

Example 6.2, Raktoe's saturated main effect plan
Raktoe [1966] showed the following saturated non-orthogonal main effect plan in a $2^{s}$ factorial:

$$
\begin{aligned}
& 0000 \\
& 0111 \\
& 1011 \\
& 1101 \\
& 1110
\end{aligned}
$$

We cannot construct the above saturated main effect plan by using the method of example 6.1. If we use the general procedure, however, we could find the above Raktoe's saturated main effect plan.

In this case,

$$
\begin{gathered}
B^{* \prime}=\left(B_{5}^{* \prime} \stackrel{\vdots}{1} B_{16-5}^{* \prime}\right)=(M, A, B, C, D, A B, A C, B C, \\
A D, B D, C D, A B C, A B D, A C D, B C D, A B C D), \\
X_{11}=\left[\begin{array}{rrrrr}
1 & -1 & -1 & -1 & -1 \\
1 & -1 & 1 & 1 & 1 \\
1 & 1 & -1 & 1 & 1 \\
1 & 1 & 1 & -1 & 1 \\
1 & 1 & 1 & 1 & -1
\end{array}\right],
\end{gathered}
$$

and

$$
X_{12}=\left[\begin{array}{rrrrrrrrrrr}
1 & 1 & 1 & 1 & 1 & 1 & -1 & -1 & -1 & -1 & 1 \\
-1 & -1 & 1 & -1 & 1 & 1 & -1 & -1 & -1 & 1 & -1 \\
-1 & 1 & -1 & 1 & -1 & 1 & -1 & -1 & 1 & -1 & -1 \\
1 & -1 & -1 & 1 & 1 & -1 & -1 & 1 & -1 & -1 & -1 \\
1 & 1 & 1 & -1 & -1 & -1 & 1 & -1 & -1 & -1 & -1
\end{array}\right]
$$

```
Using (3.6) or (3.7)
```


$\left[\begin{array}{c}\hat{A B} \\ \widehat{A C} \\ \widehat{B C} \\ \hat{A D} \\ \hat{B C} \\ \hat{C D} \\ \hat{A B C} \\ \hat{A B D} \\ \hat{A C D} \\ \hat{B C D} \\ \hat{A B C D}\end{array}\right]=\frac{1}{6}\left[\begin{array}{rrrrr}2 & 1 & 1 & 1 & 1 \\ -1 & -2 & 1 & 1 & 1 \\ -1 & 1 & -2 & 1 & 1 \\ -1 & 1 & 1 & -2 & 1 \\ -1 & 1 & 1 & 1 & -2\end{array}\right]\left[\begin{array}{l}y_{0000} \\ y_{0111} \\ y_{1011} \\ y_{1101} \\ y_{1110}\end{array}\right]$

Example 6.3, saturated main effect plans and resolution IV plans having the minimum number of treatments in a $2^{5}$ factorial
(a) Consider the defining contrasts $I \doteq A C E \doteq A B C D \doteq B D E$ in a $2^{5}$ factorial, then the alias scheme is as follows:
$\mathrm{M} \doteq \mathrm{ACE} \doteq \mathrm{ABCD} \doteq \mathrm{BDE}$
$\mathrm{E} \doteq \mathrm{AC} \doteq \mathrm{ABCDE} \simeq \mathrm{BD}$
$\mathrm{D} \doteq \mathrm{ACDE} \doteq \mathrm{ABC} \doteq \mathrm{BE}$
$\mathrm{DE} \doteq \mathrm{ACD} \doteq \mathrm{ABCE} \doteq \mathrm{B}$
$\mathrm{C} \doteq \mathrm{AE} \doteq \mathrm{ABD} \doteq \mathrm{BCDE}$
$\mathrm{CE} \doteq \mathrm{A} \doteq \mathrm{ABDE} \doteq \mathrm{BCD}$
$\mathrm{CD} \doteq \mathrm{ADE} \doteq \mathrm{AB} \doteq \mathrm{BCE}$
$\mathrm{CDE} \doteq \mathrm{AD} \doteq \mathrm{ABE} \doteq \mathrm{BC}$

After the procedure of changing row order such that the first set is $\left\{(A C E)_{0},(A B C D)_{0}\right\}$, the second set is $\left\{(A C E)_{0},(A B C D)_{1}\right\}$, the third set is $\left\{(\mathrm{ACE})_{1},(\mathrm{ABCD})_{0}\right\}$, and the fourth set is $\left\{(\mathrm{ACE})_{1},(A B C D)_{1}\right\}$ in the $\mathrm{L}_{32}$, then we obtain the following matrix:

$$
X=\left[\begin{array}{llll}
X_{11}^{*} & X_{12}^{*} & X_{13}^{*} & x_{14}^{*}  \tag{6.20}\\
X_{11}^{*} & X_{22}^{*} & X_{23}^{*} & X_{24}^{*} \\
X_{11}^{*} & X_{32}^{*} & X_{33}^{*} & X_{34}^{*} \\
X_{11}^{*} & X_{42}^{*} & X_{43}^{*} & X_{44}^{*}
\end{array}\right]
$$

where the treatment order is 00000, 00111, 01010, 01101, 10011, 10100, 11001, 11110; 01000, 01111, 00010, 00101, 11011, 11100, 10001, 10110; 11000, 11111, 10010, 10101, 01011, 01100, 00001, 00111; 10000, 10111, 11010, 11101, 00100, 00101, 01001, and 01110, and the parameter order is the same as that of $\mathrm{L}_{32^{\circ}}$

Using the method of the example 6.1 and considering $X_{11}$, we will obtain various saturated main effect plans in a $2^{5}$ factorial.
(b) Consider the defining contrast $I=A B C D$ in a $2^{5}$ factorial, then the alias scheme is as follows:

$$
\begin{aligned}
& M \doteq A B C D, E \doteq A B C D E, D \doteq A E C, D E \doteq A B C E, C \doteq A B D \\
& C E \doteq A B D E, C D \doteq A B, C D E \doteq A B E, B \doteq A C D, B E \doteq A C D E \\
& B D \doteq A C, B D E \doteq A C E, B C \doteq A D, B C E \doteq A D E, B C D \doteq A \\
& B C D E \doteq A E,
\end{aligned}
$$

where the parameter order of the left side of the equal sign is followed by the parameter order in $I_{32}$.

In the above alias scheme, the three 3 -factor interactions CDE, $B D E$, and $B C E$ are not confounded with main effects or with 2-factor interactions.

If we rearrange the treatment order with the defining contrast $I=A B C D$ and the parameter order such that the treatment order is $00000,00001,00110,00111$, 01010, 01011, 01100, 01101, 10010, 10011, 10100, 10101, 11000, 11001, 11110, 11111; 10000, 10001, 10110, 10111, 11010, 11011, 11100, 11101, 00010, 00011, 00100, 00101, 01000, 01001, 01110, and 01111, and the parameter order is $M, E, D, D E, C, C E, C D$,
$B, B E, B D, B C, B C D, B C D E, C D E, B D E, B C E ; A B C D, A B C D E, A B C, A B C E, A B D, A B D E, A B$, $A C D, A C D E, A C, A D, A, A E, A B E, A C E$, and $A D E$, then the design matrix for a $2^{5}$ factorial will be

$$
X=\left[\begin{array}{ll}
X_{11}^{*} & X_{11}^{*}  \tag{6.21}\\
X_{11}^{*} & -X_{11}^{*}
\end{array}\right]
$$

Using the method of the example 6.1, we can construct a saturated plan for the parameters $M, E, D, D E, C, C E, C D, B, B E, B D, E C, B C D$, and $B C D E \doteq A$, and this plan will be a resolution IV plan.

Let $U_{12}$ be the $16 \times 3$ matrix which consists of the last three columns of the $X_{11}^{*}$, i.e.,

$$
U_{12}^{\prime}=\left[\begin{array}{rrrrrrrrrrrrrrrr}
-1 & 1 & -1 & 1 & 1 & -1 & 1 & -1 & 1 & -1 & 1 & -1 & -1 & 1 & -1 & 1  \tag{6.22}\\
-1 & 1 & 1 & -1 & -1 & 1 & 1 & -1 & 1 & -1 & -1 & 1 & 1 & -1 & -1 & 1 \\
-1 & 1 & 1 & -1 & 1 & -1 & -1 & 1 & -1 & 1 & 1 & -1 & 1 & -1 & -1 & 1
\end{array}\right] \text {, }
$$

then we can easily find a nonsingular $3 \times 3$ matrix from the $U_{12}^{\prime}$, for example, the $13^{\text {th }}, 14^{\text {th }}$, and $16^{\text {th }}$ columns of $U_{12}^{\prime}$ make a nonsingular $3 \times 3$ matrix, then the corresponding treatments to the remaining columns of $U_{12}^{\prime}$ are 00000, 00001, 00110, 00111, 01010, 01011, 01100, 01101, 10010, 10011, 10100, 10101, and 11110. These 13 treatments will give us the minimum number of treatments for a resolution IV plan in a $2^{5}$ factorial. The method of the example 6.1 can be applied to get various plans.

Example 6.4, saturated main effect plans in a $3^{3}$ factorial
In a $3^{3}$ factorial, after the procedure of rearranging row order with defining contrast $I=A B C^{2}$, we will get the following matrix:

$$
X^{*}=\left[\begin{array}{lll}
x_{11}^{*} & x_{21}^{*} & x_{31}^{*}  \tag{6.23}\\
x_{11}^{*} & x_{22}^{*} & x_{32}^{*} \\
x_{11}^{*} & x_{23}^{*} & x_{33}^{*}
\end{array}\right]
$$

where each $X_{i j}^{*}$ is a $9 \times 9$ square matrix and the treatment order is $000,011,022$, 1011, 112, 120, 202, 210, 221; 100, 111, 122, 201, 212, 220, 002, 010, 021; 200, 211, 222, 001, 012, 020, 102, 110, and 121, and the parameter order is $M, C_{L}, C_{Q}$, $B_{L}, B_{L} C_{L}, B_{L} C_{Q}, B_{Q}, B_{Q} C_{L}, B_{Q} C_{Q} ; A_{L}, A_{L} C_{I}, A_{L} C_{Q}, A_{L} B_{L}, A_{L} B_{L} C_{L}, A_{L} B_{L} C_{Q}, A_{L} B_{Q}, A_{L} B_{Q} C_{L}$, $A_{L} B_{Q} C_{Q} ; A_{Q}, A_{Q} C_{I}, A_{Q} C_{Q}, A_{Q} B_{L}, A_{Q} B_{L} C_{L}, A_{Q} B_{L} C_{Q}, A_{Q} B_{Q}, A_{Q} B_{Q} C_{L}$, and $A_{Q} B_{Q} C_{Q}$, but we could not obtain a solution such as (6.14), because the effects $B_{L} C_{I}, B_{L} C_{Q}, B_{Q} C_{I}$, and $B_{Q} C_{Q}$ are confounded with both main effects $A_{L}$ and $A_{Q}$, respectively, i.e.,

$$
\begin{aligned}
& B_{L} C_{L} \doteq-\frac{1}{3} A_{L}=\frac{1}{3} A_{Q} \\
& B_{L} C_{Q}=-\frac{1}{3} A_{L}=-A_{Q} \\
& B_{Q} C_{L}=\frac{1}{3} A_{L}=A_{Q} \\
& B_{Q} C_{Q}=-A_{L}=A_{Q} .
\end{aligned}
$$

However, we will find that each $X_{i j}^{*}$ is a nonsingular matrix and if we rearrange the column order such that $M, A_{L}, A_{Q}, B_{L}, B_{Q}, C_{L}, C_{Q}, B_{L} C_{L}, B_{L} C_{Q}, \cdots$, and let the first $9 \times 9$ matrix of the rearranged matrix be $A_{11}$, then

$$
A_{11}=\left[\begin{array}{rrrrrrrrr}
M & A_{I} & A_{Q} & B_{I} & B_{Q} & C_{I} & C_{Q} & B_{I} C_{L} & B_{L} C_{Q}  \tag{6.24}\\
1 & -1 & 1 & -1 & 1 & -1 & 1 & 1 & -1 \\
1 & -1 & 1 & 0 & -2 & 0 & -2 & 0 & 0 \\
1 & -1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 \\
1 & 0 & -2 & -1 & 1 & 0 & -2 & 0 & 2 \\
1 & 0 & -2 & 0 & -2 & 1 & 1 & 0 & 0 \\
1 & 0 & -2 & 1 & 1 & -1 & 1 & -1 & -1 \\
1 & 1 & 1 & -1 & 1 & 1 & 1 & -1 & -1 \\
1 & 1 & 1 & 0 & -2 & -1 & 1 & 0 & 0 \\
1 & 1 & 1 & 1 & 1 & 0 & -2 & 0 & -2
\end{array}\right]
$$

 each corresponding column vectors respectively, then we can easily recognize that
 $\underline{C}_{Q},{ }_{B_{L}} C_{L}$, and ${ }_{B_{L}} C_{Q}$ are orthogonal to each other. Hence, we can say that the matrix $A_{1 I}$ is nonsingular, and then, we can make $B_{I} C_{L}$ and $B_{I} C_{Q}$ orthogonal vectors with the first 7 column vectors.

Let such new vectors of $B_{I} C_{I}$ and $B_{I} C_{Q}$ be $\underline{Z}_{I}$ and $\underline{Z}_{2}$ respectively, then by using the Schmidt method of orthogonalizing the columns we obtain

$$
\underline{z}_{1}=\underline{B}_{L} C_{L}-\frac{\left(\bar{B}_{L} C_{L} \cdot A_{L}\right)}{\left\|\underline{A}_{L}\right\|^{2}} \underline{A}_{L}-\frac{\left(B_{L} C_{L} \cdot A_{Q}\right)}{\left\|\underline{A}_{Q}\right\|^{2}} \underline{A}_{Q}=\frac{1}{3}\left[\begin{array}{r}
1  \tag{6.25}\\
-2 \\
1 \\
1 \\
1 \\
-2 \\
-2 \\
1 \\
1
\end{array}\right]
$$

and
then

$$
\left[\begin{array}{ll}
\underline{Z}_{1} & \underline{Z}_{2}
\end{array}\right]=\left[\begin{array}{rr}
1 & -1  \tag{6.27}\\
-2 & 0 \\
1 & 1 \\
1 & 1 \\
1 & -1 \\
-2 & 0 \\
-2 & 0 \\
1 & 1 \\
1 & -1
\end{array}\right] \text { ignoring the common factor. }
$$

Now, if we find a nonsingular $2 \times 2$ matrix from the $9 \times 2$ matrix, then we can construct a corresponding information matrix $X_{11}$ for saturated main effect plans.

Consider partition matrix $X_{27 \times 27}$

$$
x=\left[\begin{array}{l}
x_{1} \\
x_{2}
\end{array}\right]\left[\begin{array}{ll}
\mathrm{x}_{11} & \mathrm{x}_{12} \\
\mathrm{x}_{21} & \mathrm{x}_{22}
\end{array}\right]
$$

where $X_{11}$ is $p \times p(p<9), X_{12}$ and $X_{21}^{\prime}$ are $p \times(27-p)$ each, $X_{22}$ is $(27-p) \times(27-p)$.

Now, consider the following fraction of a $3^{3}$ factorial

$$
\begin{equation*}
Y_{p}=X_{1} B+e_{p}, p<9 \tag{6.28}
\end{equation*}
$$

where $Y_{p}^{\prime}=(000,011,022,101,112,120,202)$, then from (3.7)

$$
\begin{equation*}
\hat{B}_{p}+X_{11}^{-1} X_{12} \hat{B}_{27-p}=X_{11}^{-1} Y_{p} \tag{6.29}
\end{equation*}
$$

Now, let

$$
A_{11}^{*}=\left[\begin{array}{rrrrrrrr}
M & A_{I} & A_{Q} & B_{I} & B_{Q} & C_{I} & Z_{1} & Z_{2} \\
1 & -1 & 1 & -1 & 1 & -1 & 1 & -1 \\
1 & -1 & 1 & 0 & -2 & 0 & -2 & 0 \\
1 & -1 & 1 & 1 & 1 & 1 & 1 & 1 \\
1 & 0 & -2 & -1 & 1 & 0 & 1 & 1 \\
1 & 0 & -2 & 0 & -2 & 1 & 1 & -1 \\
1 & 0 & -2 & 1 & 1 & -1 & -2 & 0 \\
1 & 1 & 1 & -1 & 1 & 1 & -2 & 0 \\
1 & 1 & 1 & 0 & -2 & -1 & 1 & 1 \\
1 & 1 & 1 & 1 & 1 & 0 & 1 & -1
\end{array}\right]=\left[\begin{array}{lll}
X_{11} & Z_{12} \\
X_{1121} & Z_{22}
\end{array}\right],
$$

then $A_{11}^{* \prime} A_{11}^{*}$ is diagonal and $Z_{22}=\left[\begin{array}{cc}1 & 1 \\ 1 & -1\end{array}\right]$ is nonsingular. Hence,

$$
\begin{equation*}
X_{11}^{-1}=\left(U_{11}^{\prime} U_{11}\right)^{-I_{X}} X_{11}^{\prime}\left(I+\mu \mu^{\prime}\right) \tag{6.30}
\end{equation*}
$$

where

$$
\begin{align*}
& U_{11}=\left[\begin{array}{l}
x_{11} \\
X_{1121}
\end{array}\right],  \tag{6.31}\\
& \mu=-Z_{12} Z_{22}^{-1} \tag{6.32}
\end{align*}
$$

then (6.29) becomes as follows:

$$
\begin{array}{r}
\hat{B}_{p}+\left(U_{11}^{\prime} U_{11}\right)^{-1} X_{11}^{\prime}\left(I+\mu \mu^{\prime}\right) \mathrm{X}_{12} \hat{B}_{27-p} \\
 \tag{6.33}\\
=\left(U_{11}^{\prime} U_{11}\right)^{-1} X_{11}\left(I+\mu \mu^{\prime}\right) Y_{p}
\end{array}
$$

The following 26 saturated main effect plans are constructed from the set $\left\{\left(A B C^{2}\right)_{0}\right\}$ in a $3^{3}$ factorial:

| (1) 022 | (2) | 011 | (3) | 011 | (4) |  | (5) | 011 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 101 |  | 101 |  | 022 |  | 022 |  | 022 |
| 112 |  | 112 |  | 101 |  | 101 |  | 101 |
| 120 |  | 120 |  | 112 |  | 112 |  | 112 |
| 202 |  | 202 |  | 202 |  | 120 |  | 120 |
| 210 |  | 210 |  | 210 |  | 210 |  | 202 |
| 221 |  | 221 |  | 221 |  | 221 |  | 221 |
| (6) 000 | (7) | 000 | (8) | 000 | (9) | 000 | (10) | 000 |
| 101 |  | 022 |  | 022 |  | 022 |  | 022 |
| 112 |  | 112 |  | 101 |  | 101 |  | 101 |
| 120 |  | 120 |  | 120 |  | 112 |  | 112 |
| 202 |  | 202 |  | 202 |  | 120 |  | 120 |
| 210 |  | 210 |  | 210 |  | 202 |  | 202 |
| 221 |  | 221 |  | 221 |  | 221 |  | 210 |
| (11) 000 | (12) | 000 | (13) | 000 | (14) | 000 | (15) | 000 |
| 011 |  | 011 |  | 011 |  | 011 |  | 011 |
| 101 |  | 101 |  | 101 |  | 101 |  | 022 |
| 120 |  | 112 |  | 112 |  | 112 |  | 120 |
| 202 |  | 202 |  | 120 |  | 120 |  | 202 |
| 210 |  | 210 |  | 210 |  | 202 |  | 210 |
| 221 |  | 221 |  | 221 |  | 210 |  | 221 |


| (16) 000 | (17) 000 | (18) 000 | (19) 000 | (20) 000 |  |
| ---: | ---: | ---: | ---: | ---: | :--- |
| 011 | 011 | 011 | 011 | 011 |  |
| 022 | 022 | 022 | 022 | 022 |  |
| 112 | 112 | 112 | 101 | 101 |  |
| 202 | 120 | 120 | 202 | 120 |  |
| 210 | 210 | 202 | 210 | 210 |  |
| 221 | 221 | 210 | 221 | 221 |  |
| (21) 000 | $(22)$ | 000 | (23) 000 | (24) 000 | (25) 000 |
| 011 | 011 | 011 | 011 | 011 | 000 |
| 022 | 022 | 022 | 022 | 022 | 011 |
| 101 | 101 | 101 | 101 | 101 | 101 |
| 120 | 112 | 112 | 112 | 112 | 112 |
| 202 | 202 | 202 | 120 | 120 | 120 |
| 221 | 221 | 210 | 221 | 210 | 202 |

7. EFFICIENCY OF THE SATURATED MAIN EFFECT PLANS IN A $2^{n}$ FACTORIAL

There are many saturated non-orthogonal main effect plans or saturated resolution IV or V plans. Some of them are easily constructed by the method of section 6. The efficiency of each plan, however, will be different. There is no guarantee which fractions constructed by the method of example 6.1 are the most efficient plans. It will be difficult sometimes to construct the most efficient non-orthogonal main effect plans by the method of example 6.1.

We cannot expect usually that $X_{11}^{\prime} X_{11}$ is diagonal in saturated main effect plans, so our problem is to find a matrix $X_{11}$ such that (Raghavarao [1959])

$$
X_{111}^{\prime} X_{11}=\left[\begin{array}{cccccc}
r & \mu & \mu & \mu & \cdots & \mu  \tag{7.1}\\
\mu & r & \lambda & \lambda & \cdots & \lambda \\
\mu & \lambda & r & \lambda & \cdots & \lambda \\
\mu & \lambda & \lambda & r & \cdots & \lambda \\
& \vdots & & & \\
\mu & \lambda & \lambda & \lambda & \cdots & r
\end{array}\right]
$$

Now consider the condition needed to obtain a matrix of the form (7.1).
Consider the following treatment design arrays:

If we consider only the 0 level, say $O_{1}, O_{2}, O_{3}, O_{4}$ to differentiate for each factor, the above treatment design arrays have complete combinations of two 0 levels, i.e., $O_{1}, O_{2} ; O_{1}, O_{3} ; O_{1}, O_{4} ; O_{2}, O_{3} ; O_{2}, O_{4} ;$ and $O_{3}, O_{4}$, and it will be called a balanced treatment design with respect to two 0 levels.

0001
0010
0100
1000
or
1110
1101
1011
0111
are also balanced treatment designs with respect to three 0 levels or one 0 levèl, respectively.

Let an $r \times p$ matrix $A$ be the balanced treatment design arrays in a $2^{n}$ factorial, then each column in A will have the same number of elements of 0 and any $r \times 2$ submatrix in $A$ will have the same number of row vectors (0.1) or (1.0).

Consider the relationship between the treatment design arrays and corresponding design matrix $X_{11}$ of a main effect plan in a $2^{\mathrm{n}}$ factorial. In an $\mathrm{X}_{11}$, every element of the column that corresponds to the mean is $l$ and the remaining elements in the $X_{11}$ are -1 or 1 which depends on the element 0 or 1 in the treatment design arrays.

Let $\underline{x}_{0}, x_{1}, \cdots, x_{p}$ be the column vectors in an $X_{11}$, then, in a balanced main effect treatment design

$$
\begin{aligned}
\underline{x}_{0} \cdot \underline{x}_{i} & =\mu \quad \text { for } i=1,2, \cdots, p, \\
\underline{x}_{i} \cdot \underline{x}_{j} & =\lambda \quad \text { for } i \neq j \text { and } i, j=1,2, \cdots, p, \\
\underline{x}_{i}^{2} \quad & =r \quad \text { for } i=0,1, \cdots, p,
\end{aligned}
$$

where

$$
\begin{aligned}
\mu= & r-\text { number of elements which equal } 0 \text { in a column of } \mathrm{A}, \\
\lambda= & \mu \text { - number of row vectors ( } 0.1 \text { ) or (1.0) in an } r \times 2 \text { sub- } \\
& \text { matrix in } A .
\end{aligned}
$$

Hence, we can say that the balanced main effect treatment designs such as (7.2), (7.3), and (7.4) will have the information matrices of the form (7.1).

If we add the treatments 0000 or 1111 (where each factor has all the same levels) to a balanced treatment design, we will also obtain an information matrix of the form (7.1).

The following examples are the most efficient saturated non-orthogonal main effect plans in a $2^{n}$ factorial:
(1)
$n=3$
$\begin{array}{lll}0 & 0 & 0 \\ 0 & 1 & 1 \\ 1 & 0 & 1 \\ 1 & 1 & 0\end{array}$

$$
X_{11}^{\prime} X_{11}=\left[\begin{array}{llll}
4 & 0 & 0 & 0 \\
0 & 4 & 0 & 0 \\
0 & 0 & 4 & 0 \\
0 & 0 & 0 & 4
\end{array}\right]
$$

(2) $n=4$

| 0 | 0 | 0 | 0 |
| :--- | :--- | :--- | :--- |
| 0 | 1 | 1 | 1 |
| 1 | 0 | 1 | 1 |
| 1 | 1 | 0 | 1 |
| 1 | 1 | 1 | 0 |\(\quad X_{11}^{\prime} X_{11}=\left[\begin{array}{lllll}5 \& 1 \& 1 \& 1 \& 1 <br>

1 \& 5 \& 1 \& 1 \& 1 <br>
1 \& 1 \& 5 \& 1 \& 1 <br>
1 \& 1 \& 1 \& 5 \& 1 <br>
1 \& 1 \& 1 \& 1 \& 5\end{array}\right]\)
or

| 1 | 1 | 1 | 1 |
| :--- | :--- | :--- | :--- |
| 1 | 0 | 0 | 0 |
| 0 | 1 | 0 | 0 |
| 0 | 0 | 1 | 0 |
| 0 | 0 | 0 | 1 |\(\quad X_{11}^{\prime} X_{11}=\left[\begin{array}{rrrrr}5 \& -1 \& -1 \& -1 \& -1 <br>

-1 \& 5 \& 1 \& 1 \& 1 <br>
-1 \& 1 \& 5 \& 1 \& 1 <br>
-1 \& 1 \& 1 \& 5 \& 1 <br>
-1 \& 1 \& 1 \& 1 \& 5\end{array}\right]\).
(3)

$$
\begin{aligned}
& n=5 \\
& 00000 \\
& \text { - 1111 } \\
& \text { 10111 } \\
& 11011 \quad \mathrm{X}_{11}^{\prime} \mathrm{X}_{11}= \\
& \text { 11101 } \\
& 11110 \\
& X_{11}^{\prime} X_{11}=\left[\begin{array}{llllll}
6 & 2 & 2 & 2 & 2 & 2 \\
2 & 6 & 2 & 2 & 2 & 2 \\
2 & 2 & 6 & 2 & 2 & 2 \\
2 & 2 & 2 & 6 & 2 & 2 \\
2 & 2 & 2 & 2 & 6 & 2 \\
2 & 2 & 2 & 2 & 2 & 6
\end{array}\right]
\end{aligned}
$$

or

$$
\begin{array}{lllll}
1 & 1 & 1 & 1 & 1 \\
1 & 0 & 0 & 0 & 0 \\
0 & 1 & 0 & 0 & 0 \\
0 & 0 & 1 & 0 & 0 \\
0 & 0 & 0 & 1 & 0 \\
0 & 0 & 0 & 0 & 1
\end{array} \quad \mathrm{X}_{11}^{\prime} X_{11}=\left[\begin{array}{rrrrrr}
6 & -2 & -2 & -2 & -2 & -2 \\
-2 & 6 & 2 & 2 & 2 & 2 \\
-2 & 2 & 6 & 2 & 2 & 2 \\
-2 & 2 & 2 & 6 & 2 & 2 \\
-2 & 2 & 2 & 2 & 6 & 2 \\
-2 & 2 & 2 & 2 & 2 & 6
\end{array}\right]
$$

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