

A METHOD OF ANALYSIS FOR UNEQUAL NUMBERS OF REPLICATES
IN A FACTORIAL EXPERIMENT

BU-221-M

U. B. Paik

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Abstract

A method of analysis for unequal numbers of replicates in a factorial experiment is presented using Searle's [1966] method. A numerical example is included to compare with the results of Federer and Zelen [1966].

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In a $4 \times 3 \times 2$ factorial experiment in which the factors are A consisting of four levels, B of three levels and C of two levels, and if we want to test

$$\begin{aligned} H_0: A_L = 0 & \quad , \quad H_a: A_L \neq 0 \\ H_0: A_Q = 0 & \quad , \quad H_a: A_Q \neq 0 \\ & \quad \cdot \\ & \quad \cdot \\ H_0: A_C B_Q C = 0 & \quad , \quad H_a: A_C B_Q C \neq 0 \end{aligned}$$

or

$$\begin{aligned} H_0: A_L = A_Q = A_C = 0 & \quad , \quad H_a: \text{not } H_0 \\ H_0: B_L = B_Q = 0 & \quad , \quad H_a: \text{not } H_0 \\ & \quad \cdot \\ & \quad \cdot \\ H_0: A_L B_L C = A_L B_Q C = A_Q B_L C = A_Q B_Q C = A_C B_L C = A_C B_Q C = 0 & \quad , \quad H_a: \text{not } H_0 \end{aligned}$$

then we will be able to use the following method.

Let

$$Y_{ijkh} = n_{ijkh} (\mu + \tau_{ijk} + \epsilon_{ijkh}) \tag{1}$$

where

$$E(Y_{ijkh}) = \mu + \tau_{ijk}$$

$$E(\epsilon\epsilon') = \sigma^2 I$$

and

$$i = 0, 1, 2, 3; \quad j = 0, 1, 2; \quad k = 0, 1; \quad h = 0, 1, \dots, (n_{ijk} - 1)$$

$n_{ijkh} = 1$ if the treatment (ijk) is replicated at the h^{th} time
and otherwise zero.

Let X be the design matrix,

then

$$X'X = \begin{bmatrix} n_{...} & n_{000\cdot} & n_{001\cdot} & \dots & n_{321\cdot} \\ n_{000\cdot} & n_{000\cdot} & 0 & \dots & 0 \\ n_{001\cdot} & 0 & n_{001\cdot} & \dots & 0 \\ \vdots & \vdots & \vdots & & \vdots \\ n_{321\cdot} & 0 & 0 & \dots & n_{321\cdot} \end{bmatrix} \quad (2)$$

Generalized inverse of $X'X$ is

$$G = (X'X)^- = \begin{bmatrix} 0 & 0 & 0 & \dots & 0 \\ 0 & \frac{1}{n_{000\cdot}} & 0 & \dots & 0 \\ 0 & 0 & \frac{1}{n_{001\cdot}} & \dots & 0 \\ \vdots & \vdots & \vdots & & \vdots \\ 0 & 0 & 0 & \dots & \frac{1}{n_{321\cdot}} \end{bmatrix} \quad (3)$$

for which $H = GX'X =$

$$\begin{bmatrix} 0 & 0 & 0 & 0 & \dots & 0 \\ 1 & 1 & 0 & 0 & \dots & 0 \\ 1 & 0 & 1 & 0 & \dots & 0 \\ 1 & 0 & 0 & 1 & \dots & 0 \\ \vdots & \vdots & \vdots & \vdots & & \vdots \\ 1 & 0 & 0 & 0 & \dots & 1 \end{bmatrix} \quad (4)$$

Then

$$\hat{\tau} = GX'Y + (H - I)Z \tag{5}$$

where Z is arbitrary.

Let Z = 0, then we have

$$O\hat{\tau} = \begin{bmatrix} O\hat{\mu} \\ O\hat{\tau}_{000} \\ O\hat{\tau}_{001} \\ O\hat{\tau}_{010} \\ \vdots \\ O\hat{\tau}_{321} \end{bmatrix} = \begin{bmatrix} 0 \\ Y_{000.}/n_{000.} \\ Y_{001.}/n_{001.} \\ Y_{010.}/n_{010.} \\ \vdots \\ Y_{321.}/n_{321.} \end{bmatrix} \tag{6}$$

Let's consider the following linear comparisons:

A_L A_Q A_C B_L B_Q C $A_{LL}B_L$ $A_{LQ}B_Q$ $A_{QL}B_L$ $A_{QQ}B_Q$ $A_{CL}B_L$ $A_{CQ}B_Q$ A_LC A_QC $A_C C$ $B_L C$ $B_Q C$ $A_{LL}B_C$ $A_{LQ}B_C$ $A_{QL}B_C$ $A_{QQ}B_C$ $A_{CL}B_C$ $A_{CQ}B_C$	=	<table border="0"> <tr><td>0</td><td>-3</td><td>-3</td><td>-3</td><td>-3</td><td>-3</td><td>-3</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>1</td><td>1</td><td>1</td><td>1</td><td>1</td><td>1</td><td>3</td><td>3</td><td>3</td><td>3</td><td>3</td><td>3</td></tr> <tr><td>0</td><td>1</td><td>1</td><td>1</td><td>1</td><td>1</td><td>1</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>-1</td><td>1</td><td>1</td><td>1</td><td>1</td><td>1</td><td>1</td></tr> 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(7)

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We can rewrite (7) as follows:

$$\underline{b} = L' \underline{\tau} \tag{8}$$

Clearly

$$L'H = L' \tag{9}$$

then the following linear combinations are estimable

$$L' \underline{\tau} \tag{10}$$

and their best unbiased linear estimators are

$$\underline{\hat{b}} = L' \hat{\underline{\tau}}; \tag{11}$$

$$\text{var}[L' \hat{\underline{\tau}}] = L' G L \sigma^2 . \tag{12}$$

Now, using the method used by S. R. Searle [1966] in the class on linear hypotheses, the following testing of hypothesis will be possible. For example,

$$(I) \quad H_0: A_L = 0, \quad H_a: A_L \neq 0$$

$$SS_{A_L} = [Q' \hat{\underline{\tau}}]' (Q' G Q)^{-1} [Q' \hat{\underline{\tau}}] \tag{13}$$

where

$$Q' = (0 \ -3 \ -3 \ -3 \ -3 \ -3 \ -3 \ -1 \ -1 \ -1 \ -1 \ -1 \ -1 \ 1 \ 1 \ 1 \ 1 \ 1 \ 1 \ 3 \ 3 \ 3 \ 3 \ 3 \ 3)$$

and

$$F = \frac{SS_{A_L}}{\hat{\sigma}^2}, \quad \text{degrees of freedom: } (1, 19) \tag{14}$$

I conjecture that

$$E[SS_{A_L}] = (Q' G Q)^{-1} \sigma_{A_L}^2 + \sigma^2 \tag{15}$$

(II) $H_0: A_L = A_Q = A_C = 0$, $H_a: \text{not } H_0$

$$SS_A = [Q' \hat{\tau}]' (Q'GQ)^{-1} [Q' \hat{\tau}] \quad (16)$$

where

$$Q' = \begin{bmatrix} 0 & -3 & -3 & -3 & -3 & -3 & -3 & -1 & -1 & -1 & -1 & -1 & -1 & 1 & 1 & 1 & 1 & 1 & 1 & 3 & 3 & 3 & 3 & 3 & 3 \\ 0 & 1 & 1 & 1 & 1 & 1 & 1 & -1 & -1 & -1 & -1 & -1 & -1 & -1 & -1 & -1 & -1 & -1 & -1 & 1 & 1 & 1 & 1 & 1 & 1 \\ 0 & -1 & -1 & -1 & -1 & -1 & -1 & 3 & 3 & 3 & 3 & 3 & 3 & -3 & -3 & -3 & -3 & -3 & -3 & 1 & 1 & 1 & 1 & 1 & 1 \end{bmatrix}$$

and

$$F = \frac{SS_A}{3\hat{\sigma}^2} , \text{ degrees of freedom: } (3, 19) \quad (17)$$

I conjecture that

$$E[SS_A] = E \begin{bmatrix} A_L \\ A_Q \\ A_C \end{bmatrix}' (Q'GQ)^{-1} \begin{bmatrix} A_L \\ A_Q \\ A_C \end{bmatrix} + 3\sigma^2 \quad (18)$$

$$= \alpha_{11}\sigma_{A_L}^2 + \alpha_{22}\sigma_{A_Q}^2 + \alpha_{33}\sigma_{A_C}^2 + 3\sigma^2$$

where

$$(Q'GQ)^{-1} = \begin{bmatrix} \alpha_{11} & \alpha_{12} & \alpha_{13} \\ \alpha_{21} & \alpha_{22} & \alpha_{23} \\ \alpha_{31} & \alpha_{32} & \alpha_{33} \end{bmatrix}$$

The computing result above should not be different from the Federer and Zelen [1966] computing result, except for rounding errors. The latter method inverts a much smaller matrix than the method described here. The Cornell University Computing Center CUSTAT program called MINT (matrix inversion) was used to obtain the results in the following example.

Example:

	B ₀		B ₁		B ₂	
	C ₀	C ₁	C ₀	C ₁	C ₀	C ₁
A ₀	(000)	(001)	(010)	(011)	(020)	(021)
	1.11	1.52	1.09	1.27	1.21	1.24
	0.97	1.45	0.99	1.22		
A ₁	(100)	(101)	(110)	(111)	(120)	(121)
	1.30	1.55	1.03	1.24	1.12	1.27
	1.00	1.53	1.21	1.34	0.96	
A ₂	(200)	(201)	(210)	(211)	(220)	(221)
	1.22	1.38	1.34	1.40	1.34	1.46
	1.13	1.08	1.41	1.21	1.19	1.39
A ₃	(300)	(301)	(310)	(311)	(320)	(321)
	1.19	1.29	1.36	1.42	1.46	1.62
	1.03		1.16	1.39	1.03	

Analysis of Variance

Source of variation	d.f.	SS	Null hypothesis for F test
Total	43	69.3586	
Correction for mean	1	68.115684	
Among groups	23	0.936266	$A_L = A_Q = A_C = B_C = B_Q = C = \dots = A_C B_Q = 0$ or $\tau_{000} = \tau_{001} = \dots = \tau_{321}$
Within groups	19	0.306650	
A	3	0.076645	$A_L = A_Q = A_C = 0$
B	2	0.010012	$B_L = B_Q = 0$
C	1	0.369602	$C = 0$
AB	6	0.212323	$A_L B_L = A_L B_Q = A_Q B_L = A_Q B_Q = A_C B_L = A_C B_Q = 0$
AC	3	0.080971	$A_L C = A_Q C = A_C C = 0$
BC	2	0.045573	$B_L C = B_Q C = 0$
ABC	6	0.083617	$A_C B_L C = A_L B_Q C = A_Q B_L C = A_Q B_Q C = A_C B_L C = A_C B_Q C = 0$

References

Federer, W. T. and Zelen, M. [1966] "Analysis of Multifactor Classifications with Unequal Numbers of Observations." Biometrics, September.

Searle, S. R. [1966] Matrix Algebra for the Biological Sciences. John Wiley and Sons, New York. pp. 262-279.